

We shall state two theoretical results (without proof) confirming the appropriacy of causal invertible MA or AR models for stationary time series with decreasing autocorrelation function [see BD93, §3.2, §8.1].

4.27. Theorem.

Let $\{X_t\}$ be ^{an} arbitrary stationary time series with $\mu_X = 0$ and $\gamma_X(h) = \begin{cases} 0 & \text{for } h > q \\ \neq 0 & \text{for } h = q \end{cases}$. Then $\{X_t\} \sim MA(q)$, i.e. there exists a white noise $\{Z_t\} \sim WN(0, \sigma^2)$ with some variance σ^2 and parameters $\theta_1, \theta_2, \dots, \theta_q \in \mathbb{R}$, $\theta_q \neq 0$ such that:

$$X_t = Z_t + \theta_1 Z_{t-1} + \dots + \theta_q Z_{t-q}.$$

4.28. Theorem

Let $\{X_t\}$ be an arbitrary stationary time series with $\mu_X = 0$ and $\gamma_X(\cdot)$ such that $\gamma_X(h) \rightarrow 0$ for $h \rightarrow \infty$. Then for each $p \geq 0$ there exists a causal time series $Y = \{Y_t\} \sim AR(p)$ such that $\gamma_Y(h) = \gamma_X(h)$ for $h = 0, 1, \dots, p$.

4.29. Remark

- (1) It can be shown that in 4.28 there need not exist for any $q \geq 0$ an $Y = \{Y_t\} \sim MA(q)$ with analogical property, namely that $\gamma_Y(h) = \gamma_X(h)$ for $h = 0, 1, \dots, q$.
- (2) As invertible $MA(q)$ process from Theorem 4.27 may equivalently be rephrased as a causal $AR(\infty)$ -process, we can state that 4.27 is actually a special case of 4.28 which may be interpreted roughly as follows: every stationary time series with decreasing autocovariance (or autocorrelation) function may be well approximated by a causal $AR(p)$ process of sufficiently large order p making the length of mismatching parts of autocovariance functions arbitrarily small: they completely match at least at the limit $p = \infty$. As $AR(p), p < \infty$ is also trivially invertible, it is possible to show that its approximations by a causal and invertible $ARMA(p, q')$ model exist (not unique) reducing with $p' + q' < p$ the number of parameters. trailing

4.30 Theorem [cf. Theorem 3.18]

Let $\{X_t\} \sim \text{ARMA}(p, q)$: $\Phi(B)X_t = \Theta(B)Z_t$ such that $\Phi(z)$ and $\Theta(z)$ do not have common roots. Then it holds

(i) $\{X_t\}$ is causal iff $\Phi(z) \neq 0$ for all $z \in \mathbb{C}, |z| \leq 1$ (no roots inside unit disc)

In such a case ψ_j are given by $\psi(z) = \frac{\Theta(z)}{\Phi(z)} = \sum_{j=0}^{\infty} \psi_j z^j$, where $\psi := \{\psi_j\}_{j=0}^{\infty} \in \ell_1$ is unique

(ii) $\{X_t\}$ is invertible iff $\Theta(z) \neq 0$ for all $z \in \mathbb{C}, |z| \leq 1$ (no roots inside unit disc)

In such a case π_j are given by $\pi(z) = \frac{\Phi(z)}{\Theta(z)} = \sum_{j=0}^{\infty} \pi_j z^j$, where $\pi := \{\pi_j\}_{j=0}^{\infty} \in \ell_1$ is unique.

Proof:

(i) \Leftarrow : $\Phi(z) \neq 0$ at $|z| \leq 1 \Rightarrow \frac{1}{\Phi(z)}$ has no poles in $|z| \leq 1 \Rightarrow$ exists

Taylor expansion $\frac{1}{\Phi(z)} = \sum_{j=0}^{\infty} \xi_j z^j =: \xi(z)$ absolutely convergent on $|z| \leq 1$

$\Rightarrow \sum_{j=0}^{\infty} |\xi_j| < \infty$.

Then $\phi(B)X_t = \theta(B)Z_t \stackrel{4.4}{\Rightarrow} \underbrace{\xi(B)}_1 \phi(B)X_t = \underbrace{\xi(B)\theta(B)}_{=: \psi(B)} Z_t$, where

$\psi(z) = \xi(z) \cdot \theta(z) = \frac{\theta(z)}{\phi(z)}$ is absolutely convergent on $|z| \leq 1$ as well (by 3.14(6)), and $\psi \in \ell_1$ (or $\sum_{j=0}^{\infty} |\psi_j| < \infty$) by 3.14(2).

\Rightarrow : $\{X_t\}$ causal $\Rightarrow \exists \psi = \{\psi_j\} \in \ell_1$: $X_t = \sum_{j=0}^{\infty} \psi_j Z_{t-j} \stackrel{3.14(3)}{\Rightarrow} \psi(z)$ is absolutely

convergent on $|z| \leq 1$. Then in view of $\{Z_t\} \sim \text{WN}(0, \sigma^2)$, $\theta \neq 0$:

$$(*) \quad \underbrace{\theta(B)Z_t}_{\sum_{j=0}^q \theta_j Z_{t-j}} = \phi(B)X_t \stackrel{4.4}{=} \underbrace{\phi(B)\psi(B)Z_t}_{=: \eta(z) = \sum_{j=0}^{\infty} \eta_j Z_{t-j}} \quad | \langle \cdot, Z_{t-k} \rangle \neq k$$

\downarrow
on $|z| \leq 1$ by 3.14(6)

$$\sum_{j=0}^q \theta_j \underbrace{\langle Z_{t-j}, Z_{t-k} \rangle}_{\sigma^2 \delta_{j,k}} = \sum_{j=0}^{\infty} \eta_j \underbrace{\langle Z_{t-j}, Z_{t-k} \rangle}_{\sigma^2 \delta_{j,k}} \Rightarrow \begin{cases} \eta_k = \theta_k & \text{for } k=0, 1, \dots, q \\ \eta_k = 0 & \text{for } k > q \end{cases}$$

\downarrow
in view of $Z_{t-j} \perp Z_{t-k}$ for $j \neq k$

and consequently $\theta(z) = \phi(z) \cdot \psi(z)$ for $|z| \leq 1$

As $|\psi(z)| < \infty$ for $|z| \leq 1$, it must be $\phi(z) \neq 0$ for $|z| \leq 1$ [if $\phi(z_0) = 0$ for some $|z_0| \leq 1$ then $\theta(z_0) = 0$ as well, and z_0 would be a common root of $\phi(z)$ and $\theta(z)$ - contradiction]. Thus we have $\psi(z) = \frac{\theta(z)}{\phi(z)}$ for $|z| \leq 1$.

Uniqueness: $\sum_{j=0}^{\infty} \psi_j Z_{t-j} = \sum_{j=0}^{\infty} \psi'_j Z_{t-j} \quad | \langle \cdot, Z_{t-k} \rangle \Rightarrow \psi_k = \psi'_k$ as in (*).

(ii) By interchanging the role of $\phi(z)$ and $\theta(z)$. When proving the implication \Rightarrow we utilize the identities:

$$\phi(B)Z_t \stackrel{4.4}{=} \phi(B)\pi(B)X_t = \pi(B)\phi(B)X_t = \pi(B)\theta(B)Z_t \Rightarrow \phi(z) = \pi(z)\theta(z) \quad \left[\begin{array}{l} \text{as of } (*) \\ \text{as of } (*) \end{array} \right]$$

4.31 Remark

If $\{x_t\} \sim \text{ARMA}(p, q)$: $\Phi(B)x_t = \Theta(B)z_t$ where $\Phi(z)$ and $\Theta(z)$ have common roots violating the assumption of Theorem 4.30. Then a more detailed analysis shows that:

a) if all common roots lie outside the unit circle $|z|=1$, then the problem is not critical: one can still apply Theorem 4.30 after cancelling the common roots when evaluating $\psi(z) = \frac{\Theta(z)}{\Phi(z)}$ or $\pi(z) = \frac{\Phi(z)}{\Theta(z)}$

b) if one of the common roots lies on the unit circle $|z|=1$, then Theorem 4.30 remains valid with one exception: there may exist more causal or invertible solutions $\{\psi_j\}$ or $\{\pi_j\}$ (uniqueness is lost).

4.32. Computation of ψ_j (or π_j) of a causal (or invertible) ARMA(p, q)

(i) causal representation ψ_j

By 4.30.(i): $\psi(z) = \frac{\Theta(z)}{\Phi(z)} \Leftrightarrow \psi(z)\Phi(z) = \Theta(z)$

ψ_j are obtained by comparing coefficients at z^j , $j=0, 1, 2, \dots$

of the left-hand side power series product $(\psi_0 + \psi_1 z + \psi_2 z^2 + \dots)(1 - \phi_1 z - \dots - \phi_p z^p)$ with those on the right-hand side $\Theta(z) = 1 + \theta_1 z + \dots + \theta_q z^q$.

At $z^0 = 1$: $\psi_0 \cdot 1 = 1 \Rightarrow \psi_0 = 1$

$z^1 = z$: $\psi_1 \cdot 1 - \psi_0 \cdot \phi_1 = \theta_1 \Rightarrow \psi_1 = \theta_1 + \psi_0 \phi_1 = \theta_1 + \phi_1$

z^2 : $\psi_2 \cdot 1 - \psi_1 \phi_1 - \psi_0 \phi_2 = \theta_2 \Rightarrow \psi_2 = \theta_2 + (\theta_1 + \phi_1) \phi_1 + \phi_2$

where we set $\theta_j = 0$ for $j > q$ and $\phi_i = 0$ for $i > p$.

Equivalent systems of linear equations in matrix form:

(4.32a)

$$\begin{bmatrix} 1 & & & & & \\ -\phi_1 & 1 & & & & \\ \vdots & & \ddots & & & \\ -\phi_p & -\phi_{p-1} & \dots & 1 & & \\ 0 & -\phi_p & \dots & & 1 & \\ \vdots & \vdots & \vdots & & & \ddots \end{bmatrix} \begin{bmatrix} \psi_0 \\ \psi_1 \\ \psi_2 \\ \vdots \end{bmatrix} = \begin{bmatrix} 1 \\ \theta_1 \\ \theta_2 \\ \vdots \\ \theta_q \\ 0 \\ \vdots \end{bmatrix}$$

Columns: a_1, a_2, \dots

ψ Θ

(4.9a) is a SLE with lower triangular (infinite) matrix, and the above described procedure is clearly the standard solution of such systems by the so-called forward-substitution technique. Solving $\psi_0 = 1$ from the first equation in the first step we obtain from $\underline{a}_1 \psi_0 + \underline{a}_2 \psi_1 + \dots = \underline{\theta}$ after moving the first term to the right-hand side:

$\underline{a}_2 \psi_1 + \underline{a}_3 \psi_2 + \dots = \underline{\theta} - \underline{a}_1$, which is another equivalent SLE:

$$(4.9b) \begin{bmatrix} 1 & & & & & \\ -\phi_1 & 1 & & & & \\ \vdots & \vdots & \ddots & & & \\ -\phi_p & -\phi_{p-1} & & 1 & & \\ 0 & -\phi_p & & \vdots & \ddots & \\ \vdots & \vdots & & \vdots & \vdots & \ddots \end{bmatrix} \begin{bmatrix} \psi_1 \\ \psi_2 \\ \vdots \\ \vdots \\ \vdots \end{bmatrix} = \begin{bmatrix} \theta_1 + \phi_1 \\ \theta_2 + \phi_2 \\ \vdots \\ \theta_{\max(p,q)} + \phi_{\max(p,q)} \\ \vdots \end{bmatrix}$$

(ii) invertible representation π_j :

Analogously by 4.30(ii): $\pi(z) = \frac{\phi(z)}{\theta(z)} \Leftrightarrow \pi(z) \cdot \theta(z) = \phi(z)$

leads again to SLE's:

$$(4.10a) \begin{bmatrix} 1 & & & & & \\ \theta_1 & 1 & & & & \\ \vdots & \vdots & \ddots & & & \\ \theta_q & \theta_{q-1} & \dots & 1 & & \\ 0 & \theta_q & \dots & \dots & 1 & \\ \vdots & \vdots & & \vdots & \vdots & \ddots \end{bmatrix} \begin{bmatrix} \pi_0 \\ \pi_1 \\ \pi_2 \\ \vdots \\ \vdots \end{bmatrix} = \begin{bmatrix} 1 \\ -\phi_1 \\ -\phi_2 \\ \vdots \\ -\phi_p \\ 0 \\ \vdots \end{bmatrix}$$

or after substituting $\pi_0 = 1$:

$$(4.10b) \begin{bmatrix} 1 & & & & & \\ \theta_1 & 1 & & & & \\ \vdots & \vdots & \ddots & & & \\ \theta_q & \theta_{q-1} & \dots & 1 & & \\ 0 & \theta_q & \dots & \dots & 1 & \\ \vdots & \vdots & & \vdots & \vdots & \ddots \end{bmatrix} \begin{bmatrix} \pi_1 \\ \pi_2 \\ \vdots \\ \vdots \\ \vdots \end{bmatrix} = - \begin{bmatrix} \theta_1 + \phi_1 \\ \theta_2 + \phi_2 \\ \vdots \\ \theta_{\max(p,q)} + \phi_{\max(p,q)} \\ \vdots \end{bmatrix}$$

4.33. Computation of γ_x of a causal ARMA(p,q) process $X = \{X_t\}, \mu_x = 0$

A. Approximate procedure

1. Computing sufficiently many ψ_j from (4.9a) or (4.9b), $1 \leq j \leq J$
2. X causal $\Rightarrow X_t = \psi(B)z_t$ by (4.6a) $\Rightarrow \gamma_x(h) \approx \sigma^2 \sum_{j=0}^J \psi_{j+h}\psi_j, h \geq 0$

B. Exact procedure

Using the theory of homogeneous linear difference equations with constant coefficients [OD93, §3.3, 3.6] one can derive

the following formula for $\gamma_x(h)$:

(4.11) $\gamma_x(h) = \sum_{i=1}^k \sum_{j=0}^{n_i-1} \beta_{ij} h^j \xi_i^{-h}$ for $h \geq \max(p, q+1) - p = \max(0, q-p+1)$

where $\xi_i, i=1, \dots, k$, are different roots of $\phi(z)$ } $\phi(z) = -\phi_0 \prod_{i=1}^k (z - \xi_i)^{n_i}$
 $n_i \dots$ multiplicity of ξ_i , i.e.: $p = \sum_{i=1}^k n_i$ } \uparrow correction for the term at z^p to be $-\phi_0$.
 $\beta_{ij} \dots$ unknown coefficients which may be obtained from finitely many values $\psi_0 = 1, \psi_1, \dots, \psi_q$ [see below]

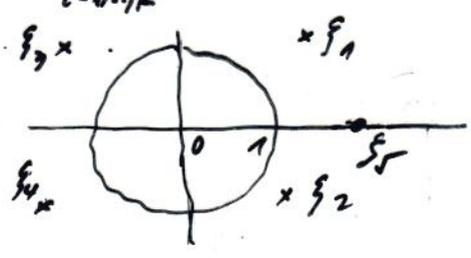
If $q-p+1 \leq 0$, then $h=0$ is in the admissible range and we get from (4.11) a formula for variance σ_x^2 :

(4.12) $\sigma_x^2 = \gamma_x(0) = \sum_{i=1}^k \beta_{i0}$ for $p > q$.

4.33.1. Remark

By 4.30(i) we have $|\xi_i| > 1$ due to causality of X . Then $|\xi_i|^{-h} = (\frac{1}{|\xi_i|})^h \rightarrow 0$ for $h \rightarrow \infty$ is faster than $h^j \rightarrow \infty \Rightarrow \gamma_x(h) \rightarrow 0$ for $h \rightarrow \infty$. The rate of decrease is exponential and accelerates with growing distance of the roots ξ_i from the unit circle, i.e. with increasing value $\min_{i=1, \dots, k} |\xi_i|$.

Observe also that ϕ_j are real iff all complex roots form complex-conjugate pairs.



4.33.2 Computation of β_{ij} and $\gamma_x(h)$, $0 \leq h < \max(p, q+1) - p$

$$\{X_t\} \sim \text{ARMA}(p, q) : \phi(B)X_t = \theta(B)Z_t, \text{ or } X_t - \sum_{i=1}^p \phi_i X_{t-i} = \sum_{j=0}^q \theta_j Z_{t-j}$$

& causa 2 : $X_t = \psi(B)Z_t$, or $X_t = \sum_{m=0}^{\infty} \psi_m Z_{t-m}$

$$X_t - \sum_{i=1}^p \phi_i X_{t-i} = \sum_{j=0}^q \theta_j Z_{t-j} \quad \left| \begin{array}{l} \times X_{t-k} \text{ for } k=0,1,2,\dots \\ \text{multiply} \end{array} \right.$$

$$X_t X_{t-k} - \sum_{i=1}^p \phi_i X_{t-i} X_{t-k} = \sum_{j=0}^q \theta_j Z_{t-j} X_{t-k} \quad \left| \begin{array}{l} E(\cdot) \dots \text{take} \\ \text{expectation and} \\ \text{subst. } X_{t-k} = \sum_{m=0}^{\infty} \psi_m Z_{t-k-m} \end{array} \right.$$

$$\underbrace{E X_t X_{t-k}}_{\gamma_x(k)} - \sum_{i=1}^p \phi_i \underbrace{E X_{t-i} X_{t-k}}_{\gamma_x(k-i)} = \sum_{j=0}^q \theta_j E \left[Z_{t-j} \sum_{m=0}^{\infty} \psi_m Z_{t-k-m} \right]$$

\Downarrow by continuity of $E(\cdot, \dots) = \langle \cdot, \dots \rangle$
with respect to the 2-nd argument ..

$$\gamma_x(k) - \sum_{i=1}^p \phi_i \gamma_x(k-i) = \sum_{j=0}^q \theta_j \sum_{m=0}^{\infty} \psi_m E Z_{t-j} Z_{t-k-m}$$

$$\underbrace{- \sum_{i=0}^p \phi_i \gamma_x(k-i)}_{\text{with } \phi_0 = -1} = \sigma^2 \sum_{j=0}^q \theta_j \psi_{j-k} \quad \begin{array}{l} \sigma^2 \delta_{j,k+m} \text{ by orthogonality } Z_{t-j} \perp Z_{t-k-m} \\ \Downarrow j=k+m \Leftrightarrow m=j-k \end{array}$$

= $\sigma^2 \sum_{j=k}^q \theta_j \psi_{j-k}$ because $\psi_{j-k} = 0$ for $j-k < 0$

having $\sum_{j=k}^q \theta_j \psi_{j-k} = 0$ for $k \geq q+1$, we arrive at the final form:

$$(4.13) \quad \underbrace{\gamma_x(k) - \sum_{i=1}^p \phi_i \gamma_x(k-i)}_{\text{LHS}} = \underbrace{\begin{cases} \sigma^2 \sum_{j=k}^q \theta_j \psi_{j-k} & \text{for } 0 \leq k < \max(p, q+1) \\ 0 & \text{for } k \geq \max(p, q+1) \end{cases}}_{\text{RHS}} \quad \text{(actually even for } k \geq q+1 \text{ if } p > q+1)$$

RHS: $0 \leq k < \max(p, q+1)$, $k \leq j \leq q \Rightarrow 0 \leq j-k \leq q-k \Rightarrow$ we substitute for $q+1$ values ψ_j , $j=0, \dots, q$ the solution of the first $q+1$ equations (4.9a) \Rightarrow RHS values are functions of θ_j and ϕ_i .

LHS: $0 \leq k < \max(p, q+1)$, $0 \leq i \leq p \Rightarrow 0-p \leq k-i \leq \max(p, q+1) - 1 - 0 < \max(p, q+1)$
 $\Rightarrow 0 \leq |k-i| \leq \max(p, q+1)$. As $\gamma_x(-h) = \gamma_x(h)$, the LHS is linear combination of $\gamma_x(h)$, $h=|k-i|$, $0 \leq h \leq \max(p, q+1)$
Substituting from (4.11) for $\gamma_x(h)$, $\max(p, q+1) - p \leq h \leq \max(p, q+1)$
we obtain $\max(p, q+1)$ linear equations for p unknowns β_{ij} and remaining $\gamma_x(h)$, $0 \leq h < \max(p, q+1) - p$, if any.

4.34. Computation of α_x of a causal ARMA(p,q) process $X = \{X_n\}, \mu_x = 0$ 37

By Theorem 4.17: $\alpha_x(n) = \phi_{nn}$ where $\phi_{nn} = \phi_{nn}^{(n)}$ is the last component of the solution to the 1-step best linear prediction problem (4.2)

If T_n is ^{nonsingular} regular then by (4.2):

$$\Phi_n = \begin{bmatrix} \phi_{nn} \\ \vdots \\ \phi_{n,n-1} \\ \phi_{nn} \end{bmatrix} = \begin{bmatrix} g_{n1} & \dots & g_{nn} \\ \vdots & & \vdots \\ \underline{g_{n-1,1}} & \dots & \underline{g_{n-1,n}} \\ \underline{g_{nn}} & \dots & \underline{g_{nn}} \end{bmatrix} \begin{bmatrix} \delta_x(1) \\ \vdots \\ \delta_x(n) \end{bmatrix} \Rightarrow \alpha_x(n) = \sum_{j=1}^n g_{nj} \delta_x(j)$$

$=: T_n^{-1}$ δ_n

4.35. Identification of ARMA type and orders based on a.c.f. and p.a.c.f. graphs.

By 4.33.1 each root ξ_i of $\phi(z)$ contributes to the sum (4.11)

by n_i (multiplicity of ξ_i) geometrically (exponentially) decreasing terms $(\sum_{j=0}^{n_i-1} \beta_{ij} h^j) \xi_i^{-h}$ for $h \geq \max(0, q-p+1)$:

polynomial factor slowing ^{down} the decrease if $n_i > 1$

Single ^{real} root \Rightarrow fastest geometrical decrease by the term $\beta_{i0} \xi_i^{-h}$

Pair of complex conjugate roots $\Rightarrow \xi_i^{-h} = |\xi_i|^{-h} (\cos(-h\alpha_i) + i \sin(-h\beta_i)) \Rightarrow$
 \Rightarrow geometrically damped sinusoidal term

\rightarrow all that mixed together \Rightarrow resulting shape of $\delta_x(h)$ for $h \geq \max(0, q-p+1)$, hereafter called U-shape

As $\alpha_x(h)$ is a partial autocorrelation to $\rho_x(h) = \frac{\delta_x(h)}{\delta_x(0)}$ a similar behaviour of $\alpha_x(h)$ is expected. Indeed, it can be shown that $\alpha_x(\cdot)$ is bounded by a U-shaped curve, for $h \geq \max(0, p-q+1)$ cf. interchange

Theorems 4.24 and 4.26 yield so-called identification point (IP):

$X \sim MA(q) \stackrel{4.24}{\Rightarrow} h_0 = q$ is IP for $\rho_x(h)$: $\rho_x(h_0) \neq 0$ and $\rho_x(h) = 0$ for $h > h_0$

$X \sim AR(p) \stackrel{4.26}{\Rightarrow} h_0 = p$ is IP for $\alpha_x(h)$: $\alpha_x(h_0) \neq 0$ and $\alpha_x(h) = 0$ for $h > h_0$

Let us summarize all the information into the following table

4.35.1. Table

	AR(p), see Fig. 4.1	MA(q), see Fig. 4.2	ARMA(p,q), see Fig. 4.3
$\rho_x(h)$	IP h_0 does not exist Shape U for $h \geq 0$	IP $h_0 = q$ (See (4.76) in Th. 4.24)	IP h_0 does not exist Shape U for $h \geq \max(0, q-p+1)$
$\alpha_x(h)$	IP $h_0 = p$ (see Theorem 4.26)	IP h_0 does not exist Bounded by U-shape for $h \geq 0$	IP h_0 does not exist Bounded by U-shape for $h \geq \max(0, p-q+1)$

4.36. Best linear prediction of zero-mean stationary time series using Innovations Algorithm. [BD 93, §5.2]

In contrast with the Durbin-Levinson algorithm 4.14, which is based on 4.11, we allow $\{X_t\}$ to be a possibly non-stationary process with mean $\mu_x = 0$ for the Innovations Algorithm (IA).

We assume again for simplicity that $t=n$ where n is the length of the prediction: $X_1, X_2, \dots, X_n, \dots$ time series up to the time $n+1$ where all $X_i \in L_2$.

Using notation analogical to Def. 4.7:

$\mathcal{L}_i := \mathcal{L}(X_1, \dots, X_i) \dots$ linear space spanned by X_1, \dots, X_i

$\mathcal{L}_0 := \{0\}$ Clearly $\mathcal{L}_0 \subset \mathcal{L}_1 \subset \dots \subset \mathcal{L}_i$

$\hat{X}_{i+1} := P_{\mathcal{L}_i}(X_{i+1}) \dots$ best linear predictions for $i=1, \dots, n$
Clearly $\hat{X}_n = 0$, which is in accordance with $E X_n = 0$ by 4.9. due to $\mu_x = 0$.

4.36.1. Theorem [Gram-Schmidt orthogonalization procedure]

$\mathcal{L}_k = \mathcal{L}(X_1 - \hat{X}_1, \dots, X_k - \hat{X}_k)$ for $k=1, \dots, n$ where $X_j - \hat{X}_j \perp X_k - \hat{X}_k$ for $j \neq k$

Proof: I. $\hat{X}_j \in \mathcal{L}_{j-1} \subset \mathcal{L}_k$, $X_j \in \mathcal{L}_j \subset \mathcal{L}_k$ for $j=1, \dots, k \Rightarrow X_j - \hat{X}_j \in \mathcal{L}_k$
for $j=1, \dots, k \Rightarrow \mathcal{L}(X_1 - \hat{X}_1, \dots, X_k - \hat{X}_k) \subset \mathcal{L}_k$

II. Denote $\mathcal{L}'_k := \mathcal{L}(X_1 - \hat{X}_1, \dots, X_k - \hat{X}_k)$. We prove by induction on k that $\mathcal{L}_k \subseteq \mathcal{L}'_k$ for $k=1, \dots, n$:

$k=1$: $\mathcal{L}_1 = \mathcal{L}(X_1) = \mathcal{L}(X_1 - \hat{X}_1) = \mathcal{L}'_1$

$k > 1$: $\mathcal{L}_{k-1} \subseteq \mathcal{L}'_{k-1} \subseteq \mathcal{L}'_k$ by induction hypothesis. Then $X_1, \dots, X_{k-1} \in \mathcal{L}_{k-1} \subseteq \mathcal{L}'_k$ and $X_k = \underbrace{X_k - \hat{X}_k}_{\in \mathcal{L}'_k} + \underbrace{\hat{X}_k}_{\in \mathcal{L}_{k-1} \subseteq \mathcal{L}'_k} \in \mathcal{L}'_k$ as well $\Rightarrow \Rightarrow X_1, \dots, X_k \in \mathcal{L}'_k \Rightarrow \mathcal{L}_k = \mathcal{L}(X_1, \dots, X_k) \subseteq \mathcal{L}'_k$.

III. Let $j \neq k$, e.g. $j < k$. Then $X_k - \hat{X}_k \perp \mathcal{L}_{k-1}$ by the orthogonal projection theorem. As $X_j - \hat{X}_j \in \mathcal{L}'_j = \mathcal{L}'_j \subseteq \mathcal{L}_{k-1}$, we have $X_k - \hat{X}_k \perp X_j - \hat{X}_j$.

4.36.2. Corollary. It holds for $k=1, \dots, n$, and $V_k := \|X_{k+1} - \hat{X}_{k+1}\|^2 = E\{X_{k+1} - \hat{X}_{k+1}\}^2$

$$\hat{X}_{k+1} = \sum_{j=1}^k \theta_{kj} (X_{k+1-j} - \hat{X}_{k+1-j}) \text{ where } V_{k-j} \theta_{kj} = \langle \hat{X}_{k+1}, X_{k+1-j} - \hat{X}_{k+1-j} \rangle = \langle X_{k+1}, X_{k+1-j} - \hat{X}_{k+1-j} \rangle$$

Proof: $\hat{X}_{k+1} \in \mathcal{L}_k = \mathcal{L}(X_1 - \hat{X}_1, \dots, X_k - \hat{X}_k) \Rightarrow \hat{X}_{k+1}$ must be a linear combination of $X_1 - \hat{X}_1, \dots, X_k - \hat{X}_k: \exists \theta_{kj}: \hat{X}_{k+1} = \sum_{j=1}^k \theta_{kj} (X_{k+1-j} - \hat{X}_{k+1-j})$

By orthogonality of $e_i := X_{k+1-i} - \hat{X}_{k+1-i}$ for $i=0, 1, \dots, k$:

$$\begin{aligned} \langle X_{k+1}, e_i \rangle &= \langle \underbrace{X_{k+1} - \hat{X}_{k+1}}_{e_0} + \hat{X}_{k+1}, e_i \rangle = \langle e_0, e_i \rangle + \langle \hat{X}_{k+1}, e_i \rangle = \\ &= \langle \sum_{j=1}^k \theta_{kj} e_j, e_i \rangle = \sum_{j=1}^k \theta_{kj} \underbrace{\langle e_j, e_i \rangle}_{V_{k-i} \delta_{ij}} = \underbrace{V_{k-i} \theta_{ki}}_{\text{for } i=1, \dots, k} \end{aligned}$$

4.36.3. Theorem. Innovations algorithm

If $\{X_n\}$ has zero mean and $\mathcal{R}(i,j) := E X_i X_j$ where the matrix $[\mathcal{R}(i,j)]_{i,j=1}^n$ is non-singular for each $n=1, 2, \dots$, then the one-step predictors $\hat{X}_{n+1}, n \geq 0$, and their mean squared errors $V_n, n \geq 1$, are given by

$$(4.15) \quad \hat{X}_{n+1} = \begin{cases} 0 & \text{if } n=0 \\ \sum_{j=1}^n \theta_{nj} (X_{n+1-j} - \hat{X}_{n+1-j}) & \text{if } n \geq 1 \end{cases}$$

and

$$\begin{aligned} V_0 &= \mathcal{R}(1,1) \\ \theta_{n,n-k} &= V_k^{-1} \left[\mathcal{R}(n+1, k+1) - \sum_{j=0}^{k-1} \theta_{k,k-j} \theta_{n,n-j} V_j \right] \\ V_n &= \mathcal{R}(n+1, n+1) - \sum_{j=0}^{n-1} \theta_{n,n-j}^2 V_j \end{aligned}$$

for $k=0, 1, \dots, n-1$ recursively in the order $V_0, \theta_{1,1}, V_1, \theta_{2,1}, \theta_{2,0}, V_2, \dots$

Proof: $v_0 = E(X_1 - \hat{X}_1)^2 = E|X_1|^2 = E X_1 X_1 = \mathcal{R}(1,1)$.

By 4.36.2 (with n instead of k and $n-k$ instead of j):

(*) $\langle X_{n+1}, X_{k+1} - \hat{X}_{k+1} \rangle = \theta_{n, n-k} v_k$ for $k=0, 1, \dots, n-1$

Nonsingularity of $[\mathcal{R}(i,j)] \Rightarrow v_k \neq 0$ [otherwise $v_k = \|X_{k+1} - \hat{X}_{k+1}\|^2 = 0$

$\Rightarrow X_{k+1} = \hat{X}_{k+1} \in \mathcal{L}(X_1, \dots, X_k) \Rightarrow \exists c_m: X_{k+1} = \sum_{m=1}^k c_m X_m \Rightarrow \mathcal{R}(k+1, j) = E(\sum_{m=1}^k c_m X_m) X_j = \sum_{m=1}^k c_m E X_m X_j = \sum_{m=1}^k c_m \mathcal{R}(m, j) \quad \forall j \Rightarrow$

$(k+1)$ -th row is linear combination of preceding rows, which contradicts nonsingularity].

Thus we can divide (*) by v_k to obtain:

$$\begin{aligned} \theta_{n, n-k} &= v_k^{-1} \langle X_{n+1}, X_{k+1} - \hat{X}_{k+1} \rangle = v_k^{-1} \{ \langle X_{n+1}, X_{k+1} \rangle - \langle X_{n+1}, \hat{X}_{k+1} \rangle \} = \\ &= v_k^{-1} \{ \mathcal{R}(n+1, k+1) - \sum_{i=1}^k \theta_{n, n-i} \langle X_{n+1}, X_{i+1} - \hat{X}_{i+1} \rangle \} \quad \begin{matrix} j \leftrightarrow k-j \\ = \dots \end{matrix} \\ &= v_k^{-1} \{ \mathcal{R}(n+1, k+1) - \sum_{i=0}^{k-1} \theta_{n, n-i} \langle X_{n+1}, X_{i+1} - \hat{X}_{i+1} \rangle \}. \end{aligned}$$

$\theta_{n, n-j} v_j$ by (*)

$$\begin{aligned} v_n &= \|X_{n+1} - \hat{X}_{n+1}\|^2 = \langle X_{n+1} - \hat{X}_{n+1}, X_{n+1} - \hat{X}_{n+1} \rangle = \\ &= \langle X_{n+1}, X_{n+1} - \hat{X}_{n+1} \rangle - \langle \hat{X}_{n+1}, X_{n+1} - \hat{X}_{n+1} \rangle = \\ &= \langle X_{n+1}, X_{n+1} \rangle - \langle X_{n+1} - \hat{X}_{n+1} + \hat{X}_{n+1}, \hat{X}_{n+1} \rangle = \mathcal{R}(n+1, n+1) - \end{aligned}$$

$$\begin{aligned} &= \langle X_{n+1} - \hat{X}_{n+1}, \hat{X}_{n+1} \rangle - \langle \hat{X}_{n+1}, \hat{X}_{n+1} \rangle = \mathcal{R}(n+1, n+1) - \mathcal{R}(n+1, n+1) \\ &= \langle X_{n+1} - \hat{X}_{n+1}, \hat{X}_{n+1} \rangle - \langle \hat{X}_{n+1}, \hat{X}_{n+1} \rangle = \mathcal{R}(n+1, n+1) - \mathcal{R}(n+1, n+1) \end{aligned}$$

$$\sum_{i=0}^{n-1} \theta_{n, n-i} \langle X_{i+1} - \hat{X}_{i+1}, \hat{X}_{n+1} \rangle = \sum_{k=0}^{n-1} \sum_{j=0}^{n-1} \theta_{n, n-k} \theta_{n, n-j} \langle X_{k+1} - \hat{X}_{k+1}, X_{j+1} - \hat{X}_{j+1} \rangle = \sum_{k=0}^{n-1} \sum_{j=0}^{n-1} \theta_{n, n-k} \theta_{n, n-j} v_k \delta_{j,k}$$

$\mathcal{R}(n+1, n+1) - \sum_{k=0}^{n-1} \theta_{n, n-k}^2 v_k$