Econometrics - Lecture 6

GMM-Estimator and Econometric Models

Contents

- GMM Estimation
- Econometric Models
- Dynamic Models
- Multi-equation Models
- Time Series Models
- Models for Limited Dependent Variables
- Panel Data Models
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From OLS to IV Estimation

Linear model $y_i = x_i^{\beta} + \varepsilon_i$

OLS estimator: solution of the K normal equations

 $1/N \Sigma_{i}(y_{i} - x_{i}^{*}b) x_{i} = 0$

Corresponding moment conditions

 $\mathsf{E}\{\varepsilon_i | x_i\} = \mathsf{E}\{(y_i - x_i;\beta) | x_i\} = 0$

 IV estimator given R instrumental variables z_i which may overlap with x_i: based on the R moment conditions

 $\mathsf{E}\{\varepsilon_i \ z_i\} = \mathsf{E}\{(y_i - x_i;\beta) \ z_i\} = 0$

 IV estimator: solution of corresponding sample moment conditions

$$\frac{1}{N}\sum_{i}(y_{i}-x_{i}^{\prime}\hat{\beta}_{IV})z_{i}=0$$

The IV Estimator

General case: R moment conditions

R = *K*: one unique solution, the IV estimator; identified model

 *β*_{IV} = (∑_i z_i x'_i)⁻¹ ∑_i z_i y_i = (Z ' X)⁻¹ Z ' y

 R > *K*: minimization of the weighted quadratic form in the sample moments

$$Q_N(\boldsymbol{\beta}) = \left[\frac{1}{N}\sum_i (y_i - x_i'\boldsymbol{\beta}) z_i\right]' W_N\left[\frac{1}{N}\sum_i (y_i - x_i'\boldsymbol{\beta}) z_i\right]$$

with a RxR positive definite weighting matrix W_N gives the generalized instrumental variable (GIV) estimator

$$\hat{\boldsymbol{\beta}}_{IV} = (X'ZW_N Z'X)^{-1} X'ZW_N Z'y$$

Optimal weighting matrix: $W_N^{opt} = [1/N(Z'Z)]^{-1}$; corresponds to the most efficient IV estimator

$$\hat{\beta}_{IV} = (X'Z(Z'Z)^{-1}Z'X)^{-1}X'Z(Z'Z)^{-1}Z'y$$

Generalized Method of Moments (GMM) Estimation

The model is characterized by *R* moment conditions and the corresponding equations

 $\mathsf{E}\{f(w_i, z_i, \theta)\} = 0$

[cf. E{ $(y_i - x_i, \beta) z_i$ } = 0]

- f(.): R-vector function
- *w*_i: vector of observable variables, exogenous or endogenous
- z_i: vector of instrumental variables
- θ: *K*-vector of unknown parameters

Sample equivalents $g_N(\theta)$ of moment conditions should fulfill

$$g_N(\theta) = \frac{1}{N} \sum_i f(w_i, z_i, \theta) = 0$$

Estimates $\hat{\theta}$ are chosen such that the sample moment conditions are fulfilled

GMM Estimation

 $R \ge K$ is a necessary condition for GMM estimation

• R = K: unique solution, the K-vector $\hat{\theta}$, of

 $g_{\rm N}(\theta) = 0$

if f(.) is nonlinear in θ , numerical solution might be derived

• R > K: in general, no choice $\hat{\theta}$ for the *K*-vector θ will result in $g_N(\hat{\theta})$ = 0 for all *R* equations; for a good choice $\hat{\theta}$, $g_N(\hat{\theta}) \sim 0$, i.e., all components of $g_N(\hat{\theta})$ are close to zero estimate $\hat{\theta}$ is obtained through minimization wrt θ of the quadratic

form

 $Q_{N}(\theta) = g_{N}(\theta)' W_{N} g_{N}(\theta)$

 $W_{\rm N}$: symmetric, positive definite weighting matrix

The GMM Estimator

Weighting matrix $W_{\rm N}$

- Different weighting matrices result in different consistent estimators with different covariance matrices
- Optimal weighting matrix

 $W_{N}^{opt} = [E\{f(w_{i}, z_{i}, \theta) f(w_{i}, z_{i}, \theta)'\}]^{-1}$

i.e., the inverse of the covariance matrix of the sample moments

For
$$R = K$$
: $W_N = I_N$ with unit matrix I_N

Minimization of $Q_N(\theta) = g_N(\theta)^{\prime} W_N g_N(\theta)$: For nonlinear f(.)

- Numerical optimization algorithms
- $W_{\rm N}$ depends on θ ; iterative optimization

Example: The Linear Model

Model: $y_i = x_i^{\beta} + \varepsilon_i$ with $E\{\varepsilon_i | x_i\} = 0$ and $V\{\varepsilon_i\} = \sigma_{\varepsilon}^2$

Moment or orthogonality conditions:

 $\mathsf{E}\{\varepsilon_{\mathsf{t}} x_{\mathsf{t}}\} = \mathsf{E}\{(y_{\mathsf{t}} - x_{\mathsf{t}} \beta)x_{\mathsf{t}}\} = 0$

 $f(.) = (y_i - x_i^{\beta})x_i, \theta = \beta$, instrumental variables: x_i ; moment conditions are exogeneity conditions for x_i

Sample moment conditions:

 $1/N \sum_{i} (y_{i} - x_{i} b) x_{i} = 1/N \sum_{i} e_{i} x_{i} = g_{N}(b) = 0$

- With $W_N = I_N$, $Q_N(\beta) = [1/N]^2 (\Sigma_i \varepsilon_i x_i)'(\Sigma_i \varepsilon_i x_i) = [1/N]^2 X' \varepsilon \varepsilon' X$
- OLS and GMM estimators coincide, give the estimator *b*, but
 - OLS: residual sum of squares $S_N(b) = 1/N \Sigma_i e_i^2$ has its minimum
 - $\Box \quad \text{GMM: } Q_{N}(b) = [1/N]^{2} (\Sigma_{i} e_{i} x_{i})'(\Sigma_{i} e_{i} x_{i}) = 0$

Linear Model, $E\{\varepsilon_t x_t\} \neq 0$

Model $y_i = x_i^{\beta} + \varepsilon_i$ with $V{\varepsilon_i} = \sigma_{\varepsilon}^2$, $E{\varepsilon_i x_i} \neq 0$ and *R* instrumental variables z_i

Moment conditions:

 $\mathsf{E}\{\varepsilon_i \ z_i\} = \mathsf{E}\{(y_i - x_i^{\,i}\beta)z_i\} = 0$

Sample moment conditions:

 $1/N \Sigma_i (y_i - x_i'b) z_i = g_N(b) = 0$

Identified case (R = K): the single solution is the IV estimator

 $b_{IV} = (Z'X)^{-1} Z'y$

• Over-identified case (R > K): GMM estimator from min_{β} Q_N(β)= min_{β} $g_N(\beta)$ ' $W_N g_N(\beta)$

Linear Model: GMM Estimator

Minimization of $Q_N(\beta) = \min_{\beta} g_N(\beta)' W_N g_N(\beta)$ wrt β :

• For $W_N = I$, the first order conditions are

$$\frac{\partial Q_N(\beta)}{\partial \beta} = 2\left(\frac{\partial g_N(\beta)}{\partial \beta}\right)' g_N(\beta) = 2\left(\frac{1}{N}X'Z\right)\left(\frac{1}{N}Z'y - \frac{1}{N}Z'X\beta\right) = 0$$

resulting in the estimator

 $b = [(X'Z)(Z'X)]^{-1} (X'Z)Z'y$

b coincides with the IV estimator if R = K

The optimal weighting matrix $W_N^{\text{opt}} = (E\{\varepsilon_i^2 z_i z_i'\})^{-1}$ is estimated by $W_N^{opt} = \left(\frac{1}{N}\sum_i e_i^2 z_i z_i'\right)^{-1}$

generalizes the covariance matrix of the GIV estimator to White's heteroskedasticity-consistent covariance matrix estimator (HCCME)

Example: Labour Demand

Verbeek's data set "labour2": Sample of 569 Belgian companies (data from 1996)

- Variables
 - labour: total employment (number of employees)
 - capital: total fixed assets
 - wage: total wage costs per employee (in 1000 EUR)
 - output: value added (in million EUR)
- Labour demand function

labour = $\beta_1 + \beta_2^*$ *output* + β_3^* *capital*

Labour Demand Function: OLS Estimation

In logarithmic transforms: Output from **GRETL**

Dependent variable : I_LABOUR Heteroskedastic-robust standard errors, variant HC0,

coefficient	std. error	t-ratio	p-value
const 3,01483 I_ OUTPUT 0,878061	0,0566474 0,0512008	53,22 17,15	1,81e-222 *** 2,12e-053 ***
I_CAPITAL 0,003699	0,0429567	0,08610	0,9314
Mean dependent var	4,488665	S.D. dependent var	1,171166
Sum squared resid	158,8931	S.E. of regression	0,529839
R- squared	0,796052	Adjusted R-squared	0,795331
F(2, 129)	768,7963	P-value (F)	4,5e-162
Log-likelihood	-444,4539	Akaike criterion	894,9078
Schwarz criterion	907,9395	Hannan-Quinn	899,9928

Specification of GMM Estimation

GRETL: Specification of function and orthogonality conditions for labour demand model

```
# initializations go here
matrix X = {const , I_OUTPUT, I_CAPITAL}
series e = 0
scalar b1 = 0
scalar b2 = 0
scalar b3 = 0
matrix V = I(3)

gmm e = I_LABOuR - b1*const - b2*I_OUTPUT - b3*I_CAPITAL
orthog e; X
weights V
params b1 b2 b3
end gmm
```

Labour Demand Function: GMM Estimation

In logarithmic transforms: Output from **GRETL**

Using numerical derivatives Tolerance = 1,81899e-012 Function evaluations: 44 Evaluations of gradient: 8

Model 8: 1-step GMM, using observations 1-569 e = I_LABOUR - b1*const - b2*I_OUTPUT - b3*I_CAPITAL

	estimate	std. error	t-ratio	p-value
b1	3,01483	0,0566474	53,22	0,0000 ***
b2	0,878061	0,0512008	17,15	6,36e-066 ***
b3	0,00369851	0,0429567	0,0861	0 0,9314

GMM criterion: Q = 1,1394e-031 (TQ = 6,48321e-029)

Labour Demand Functions: Comparison of Estimates

OLS and GMM estimates coincide

	OLS	GMM
const	3,015	3,015
	0,057	0,057
L_OUTPUT	0,878	0,878
	0,051	0,051
I_CAPITAL	0,0037	0,0037
	0,0430	0,0430

GMM Estimator: Properties

Under weak regularity conditions, the GMM estimator is

- consistent (for any W_N)
- most efficient if $W_N = W_N^{opt} = [E\{f(w_i, z_i, \theta) | f(w_i, z_i, \theta)'\}]^{-1}$
- asymptotically normal: $\sqrt{N}(\hat{\theta} \theta) \rightarrow N(0, V^{-1})$

where $V = D W_N^{\text{opt}} D$ with the KxR matrix of derivatives $D = E \left\{ \frac{\partial f(w_i, z_i, \theta)}{\partial \theta} \right\}$

The covariance matrix V^{-1} can be estimated by substituting the population parameters θ by sample equivalents $\hat{\theta}$ evaluated at the GMM estimates in *D* and W_N^{opt}

GMM Estimator: Calculation

- 1. One-step GMM estimator: Choose a positive definite W_N , *e.g.*, $W_N = I_N$, optimization gives $\hat{\theta}_1$ (consistent, but not efficient)
- 2. Two-step GMM estimator: use the one-step estimator $\hat{\theta}_1$ to estimate $V = D W_N^{opt} D'$, repeat optimization with $W_N = V^{-1}$; this gives $\hat{\theta}_2$
- 3. Iterated GMM estimator: Repeat step 2 until convergence
- If R = K, the GMM estimator is the same for any W_N , only step 1 is needed; the objective function $Q_N(\theta)$ is zero at the minimum

If R > K, step 2 is needed to achieve efficiency

GMM and Other Estimation Methods

- GMM estimation generalizes the method of moments estimation
- Allows for a general concept of moment conditions
- Moment conditions are not necessarily linear in the parameters to be estimated
- Encompasses various estimation concepts such as OLS, GLS, IV, GIV, ML

moment conditions

OLS $E\{(y_i - x_i'\beta) x_i\} = 0$ GLS $E\{(y_i - x_i'\beta) x_i/\sigma^2(x_i)\} = 0$ IV $E\{(y_i - x_i'\beta) z_i\} = 0$

ML $E\{\partial/\partial\beta f[\varepsilon_i(\beta)]\} = 0$

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Klein's Model 1

$$\begin{aligned} C_t &= \alpha_1 + \alpha_2 P_t + \alpha_3 P_{t-1} + a_4 (W_t^{p} + W_t^{g}) + \varepsilon_{t1} \quad (\text{consumption}) \\ I_t &= \beta_1 + \beta_2 P_t + \beta_3 P_{t-1} + \beta_4 K_{t-1} + \varepsilon_{t2} \quad (\text{investments}) \\ W_t^{p} &= \gamma_1 + \gamma_2 X_t + \gamma_3 X_{t-1} + \gamma_4 t + \varepsilon_{t3} \quad (\text{private wages and salaries}) \\ X_t &= C_t + I_t + G_t \\ K_t &= I_t + K_{t-1} \\ P_t &= X_t - W_t^{p} - T_t \end{aligned}$$

C (consumption), P (profits), W^p (private wages and salaries), W^g (public wages and salaries), I (investments), K (capital stock), X (production), G (governmental expenditures without wages and salaries), T (taxes) and t [time (trend)]

Endogenous: *C*, *I*, *W*^p, *X*, *P*, *K*; exogeneous: *W*^g, *G*, *T*, *t*, *P*₋₁, *K*₋₁, *X*₋₁

Early Econometric Models

Klein's Model

- Aims:
 - to forecast the development of business fluctuations and
 - to study the effects of government economic-political policy
- Successful forecasts of
 - economic upturn rather than a depression after World War II
 - mild recession at the end of the Korean War

Model	year	eq's
Tinbergen	1936	24
Klein	1950	6
Klein & Goldberger	1955	20
Brookings	1965	160
Brookings Mark II	1972	~200

Econometric Models

Basis: the multiple linear regression model

- Adaptations of the model
 - Dynamic models
 - Systems of regression models
 - Time series models
- Further developments
 - Models for panel data
 - Models for spatial data
 - Models for limited dependent variables

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Dynamic Models: Examples

Demand model: describes the quantity *Q* demanded of a product as a function of its price *P* and consumers' income *Y*

(a) Current price and current income determine the demand (static model):

 $Q_t = \beta_1 + \beta_2 P_t + \beta_3 Y_t + \varepsilon_t$

(b) Current price and income of the previous period determine the demand (dynamic model):

 $Q_t = \beta_1 + \beta_2 P_t + \beta_3 Y_{t-1} + \varepsilon_t$

(c) Current price and demand of the previous period determine the demand (autoregressive model):

$$Q_{t} = \beta_{1} + \beta_{2}P_{t} + \beta_{3}Q_{t-1} + \varepsilon_{t}$$

Dynamic of Processes

Static processes: independent variables have a direct effect, the adjustment of the dependent variable on the realized values of the independent variables is completed within the current period, the process is assumed to be always in equilibrium

Static models may be unsuitable:

- (a) Some activities are determined by the past, such as: energy consumption depends on past investments into energyconsuming systems and equipment
- (b) Actors of the economic processes often respond with delay, e.g., due to the duration of decision-making and procurement processes
- (c) Expectations: e.g., consumption depends not only on current income but also on income expectations in future; modelling of income expectation based on past income development

Elements of Dynamic Models

 Lag-structures, distributed lags: describe the delayed effects of one or more regressors on the dependent variable; e.g., the lagstructure of order s or DL(s) model (DL: distributed lag)

 $Y_t = \alpha + \sum_{i=0}^{s} \beta_i X_{t-i} + \varepsilon_t$

- 2. Geometric lag-structure, Koyck's model: infinite lag-structure with $\beta_i = \lambda_0 \lambda^i$ (0 < λ < 1)
- 3. ADL-model: autoregressive model with lag-structure, e.g., the ADL(1,1)-model

$$Y_t = \alpha + \varphi Y_{t-1} + \beta_0 X_t + \beta_1 X_{t-1} + \varepsilon_t$$

4. Error-correction model

 $\Delta Y_t = - (1-\varphi)(Y_{t-1} - \mu_0 - \mu_1 X_{t-1}) + \beta_0 \Delta X_t + \varepsilon_t$ obtained from the ADL(1,1)-model with $\mu_0 = \alpha/(1-\varphi)$ und $\mu_1 = (\beta_0 + \beta_1)/(1-\varphi)$

The Koyck Transformation

Transforms the model

 $Y_{t} = \lambda_{0} \Sigma_{i} \lambda^{i} X_{t-i} + \varepsilon_{t}$

into an autoregressive model ($v_t = \varepsilon_t - \lambda \varepsilon_{t-1}$):

 $Y_{t} = \lambda Y_{t-1} + \lambda_0 X_t + v_t$

- The model with infinite lag-structure in X becomes a model
 - with an autoregressive component λY_{t-1}
 - with a single regressor X_t and
 - with autocorrelated error terms
- Econometric applications
 - The adaptive expectations model

Example: Investments determined by expected profit X^{e} :

 $X^{e}_{t+1} = \lambda X^{e}_{t} + (1 - \lambda) X_{t} \text{ (with } 0 < \lambda < 1)$

The partial adjustment model

Example: K_{t}^{p} : planned stock for *t*; strategy for adapting K_{t} on K_{t}^{p}

$$K_{\rm t}-K_{\rm t-1}=\delta(K^{\rm p}_{\rm t}-K_{\rm t-1})$$

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Example: Income and Consumption

Consumption C_t and disposable income Y_t are simultaneously determined by

$$C_{t} = \beta_{1} + \beta_{2}Y_{t} + \varepsilon_{t} \qquad (A)$$

$$Y_{t} = C_{t} + I_{t} \qquad (B)$$

- The disposable income Y_t is determined by the consumption C_t
- Equations (A) and (B) are the structural equations or the structural form of the simultaneous equation model that describes both C_t and Y_t
- The coefficients β_1 and β_2 are behavioural parameters
- In equation (A), Y_t is endogenous: The OLS estimates b₁ and b₂ are biased and not consistent

Multi-equation Models

Economic phenomena are usually characterized by the behaviour of more than one dependent variable Multi-equation model: the number of equations determines the number of dependent variables which are described by the model

Characteristics of multi-equation models:

- Types of equations
- Types of variables
- Identifiability

Klein's Model 1

$$C_{t} = \alpha_{1} + \alpha_{2}P_{t} + \alpha_{3}P_{t-1} + a_{4}(W_{t}^{p} + W_{t}^{g}) + \varepsilon_{t1} \quad (\text{consumption})$$

$$I_{t} = \beta_{1} + \beta_{2}P_{t} + \beta_{3}P_{t-1} + \beta_{4}K_{t-1} + \varepsilon_{t2} \quad (\text{investments})$$

$$W_{t}^{p} = \gamma_{1} + \gamma_{2}X_{t} + \gamma_{3}X_{t-1} + \gamma_{4}t + \varepsilon_{t3} \quad (\text{private wages and salaries})$$

$$X_{t} = C_{t} + I_{t} + G_{t}$$

$$K_{t} = I_{t} + K_{t-1}$$

$$P_{t} = X_{t} - W_{t}^{p} - T_{t}$$

$$C \quad (\text{consumption}), P \quad (\text{profits}), W^{p} \quad (\text{private wages and salaries}), W^{g} \quad (\text{public})$$

wages and salaries), *I* (<u>investments</u>), *K* (<u>capital stock</u>), *X* (<u>production</u>), *G* (governmental expenditures without wages and salaries), *T* (taxes) and *t* [time (trend)]

Endogenous: *C*, *I*, *W*^p, *X*, *P*, *K*; exogeneous: *W*^g, *G*, *T*, *t*, *P*₋₁, *K*₋₁, *X*₋₁

Types of Equations

- Behavioural or structural equations: describe the behaviour of a dependent variable as a function of explanatory variables
- Definitional identities: define how a variable is defined as the sum of other variables, e.g., decomposition of gross domestic product as the sum of its consumption components

Example: Klein's model 1: $X_t = C_t + I_t + G_t$

 Equilibrium conditions: assume a certain relationship, which can be interpreted as an equilibrium

Example: equality of demand (Q^d) and supply (Q^s) in a market model: $Q_t^d = Q_t^s$

Definitional identities and equilibrium conditions have no error terms

Types of Variables

Specification of a multi-equation model: definition of

- variables which are explained by the model (endogenous variables)
- other variables which are used in the model

Number of equations needed in the model: same number as that of the endogenous variables in the model

Explanatory or exogenous variables: uncorrelated with error terms

- strictly exogenous variables: uncorrelated with error terms ε_{t+i} (for any $i \neq 0$)
- predetermined variables: uncorrelated with current and future error terms ($ε_{t+i}$, *i* ≥ 0); lagged explanatory variables

Error terms:

- Uncorrelated over time
- Error terms from different equations and same observation period typically correlated, contemporaneous correlation

Systems of Regression Equations

Economic processes encompass the simultaneous developments as well as interrelations of a set of dependent variables

 For modelling economic processes: system of relations, typically in the form of regression equations: multi-equation model

Example: Two dependent variables y_{t1} and y_{t2} are modelled as

 $y_{t1} = x_{t1}^{*}\beta_{1} + \varepsilon_{t1}$ $y_{t2} = x_{t2}^{*}\beta_{2} + \varepsilon_{t2}$ with V{ ε_{ti} } = σ_{i}^{2} for i = 1, 2, Cov{ $\varepsilon_{t1}, \varepsilon_{t2}$ } = $\sigma_{12} \neq 0$ Typical situations:

- 1. The set of regressors x_{t1} and x_{t2} coincide
- 2. The set of regressors x_{t1} and x_{t2} differ, may overlap
- 3. Regressors contain one or both dependent variables
- 4. Regressors contain lagged variables

Types of Multi-equation Models

Multivariate regression or multivariate multi-equation model

- A set of regression equations, each explaining one of the dependent variables
 - Possibly common explanatory variables
 - Seemingly unrelated regression (SUR) model: each equation is a valid specification of a linear regression, related to other equations only by the error terms
 - □ See cases 1 and 2 of "typical situations"

Simultaneous equations models

- Describe the relations within the system of economic variables
 - □ in form of model equations
 - See cases 3 and 4 of "typical situations"

Error terms: dependence structure is specified by means of second moments or as joint probability distribution

Capital Asset Pricing Model

Capital asset pricing (CAP) model: describes the return R_i of asset *i*

$$R_{\rm i} - R_{\rm f} = \beta_{\rm i}({\rm E}\{R_{\rm m}\} - R_{\rm f}) + \varepsilon_{\rm i}$$

with

- \square $R_{\rm f}$: return of a risk-free asset
- \square $R_{\rm m}$: return of the market's optimal portfolio
- β_i: indicates how strong fluctuations of the returns of asset *i* are determined by fluctuations of the market as a whole
- Knowledge of the return difference R_i R_f will give information on the return difference R_i - R_f of asset *j*, at least for some assets
- Analysis of a set of assets *i* = 1, ..., s
 - The error terms ε_i, *i* = 1, ..., *s*, represent common factors, e.g., inflation rate, have a common dependence structure
 - Efficient use of information: simultaneous analysis

A Model for Investment

Grunfeld investment data (Grunfeld & Griliches, 1960): Panel data set on gross investments I_{it} of firms i = 1, ..., 6 over 20 years and related data

Investment decisions are assumed to be determined by

 $I_{it} = \beta_{i1} + \beta_{i2}F_{it} + \beta_{i3}C_{it} + \varepsilon_{it}$

with

- F_{it} : market value of firm *i* at the end of year *t*-1
- C_{it} : value of stock of plant and equipment at the end of year *t*-1
- Simultaneous analysis of equations for the various firms *i*: efficient use of information
 - Error terms for the firms include common factors such as economic climate
 - Coefficients may be the same for the firms

The Hog Market

Model equations:

 $\begin{array}{l} Q^{d} = \alpha_{1} + \alpha_{2}P + \alpha_{3}Y + \varepsilon_{1} \quad (\text{demand equation}) \\ Q^{s} = \beta_{1} + \beta_{2}P + \beta_{3}Z + \varepsilon_{2} \quad (\text{supply equation}) \\ Q^{d} = Q^{s} \quad (\text{equilibrium condition}) \end{array}$ with Q^{d} : demanded quantity, Q^{s} : supplied quantity, P: price, Y: income, and Z: costs of production, or $\begin{array}{c} Q = \alpha_{1} + \alpha_{2}P + \alpha_{3}Y + \varepsilon_{1} \quad (\text{demand equation}) \\ Q = \beta_{1} + \beta_{2}P + \beta_{3}Z + \varepsilon_{2} \quad (\text{supply equation}) \end{array}$

- Model describes quantity and price of the equilibrium transactions
- Model determines simultaneously Q and P, given Y and Z
- Error terms
 - May be correlated: $Cov{\epsilon_1, \epsilon_2\} \neq 0$
 - Simultaneous analysis necessary for efficient use of information

Examples of Multi-equation Models

Multivariate regression models

- Capital asset pricing (CAP) model: for all assets, return R_i (or risk premium R_i R_f) is a function of E{R_m} R_f; dependence structure of the error terms caused by common variables
- Model for investment: firm-specific regressors, dependence structure of the error terms like in CAP model
- Seemingly unrelated regression (SUR) models

Simultaneous equations models

- Hog market model: endogenous regressors, dependence structure of error terms
- Klein's model I: endogenous regressors, dynamic model, dependence of error terms from different equations and possibly over time

Single- vs. Multi-equation Models

Complications for estimation of parameters of multi-equation models:

- Dependence structure of error terms
- Violation of exogeneity of regressors

Example: Hog market model, demand equation

 $Q = \alpha_1 + \alpha_2 P + \alpha_3 Y + \varepsilon_1$

• Covariance matrix of
$$\varepsilon = (\varepsilon_1, \varepsilon_2)$$

$$\operatorname{Cov}\{\varepsilon\} = \begin{pmatrix} \sigma_1^2 & \sigma_{12} \\ \sigma_{12} & \sigma_2^2 \end{pmatrix}$$

P is not exogenous: Cov{*P*,ε₁} = (σ₁² - σ₁₂)/(β₂ - α₂) ≠ 0
 Statistical analysis of multi-equation models requires methods adapted to these features

Multi-equation Models: Estimation of Parameters

Estimation procedures

- Multivariate regression models
 - □ FGLS , GLS, ML
- Simultaneous equations models
 - Single equation methods: indirect least squares (ILS), two stage least squares (TSLS), limited information ML (LIML)
 - System methods of estimation: three stage least squares (3SLS), full information ML (FIML)
 - Dynamic models: estimation methods for vector autoregressive (VAR) and vector error correction (VEC) models

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Types of Trend

Trend: The expected value of a process $\{Y_t, t = 1, 2, ...\}$ increases or decreases with time

Deterministic trend: a function f(t) of the time t, describing the evolution of E{Y_t} over time

 $Y_t = f(t) + \varepsilon_t, \varepsilon_t$: white noise

Example: $Y_t = \alpha + \beta t + \varepsilon_t$ describes a linear trend of *Y*; an increasing trend corresponds to $\beta > 0$

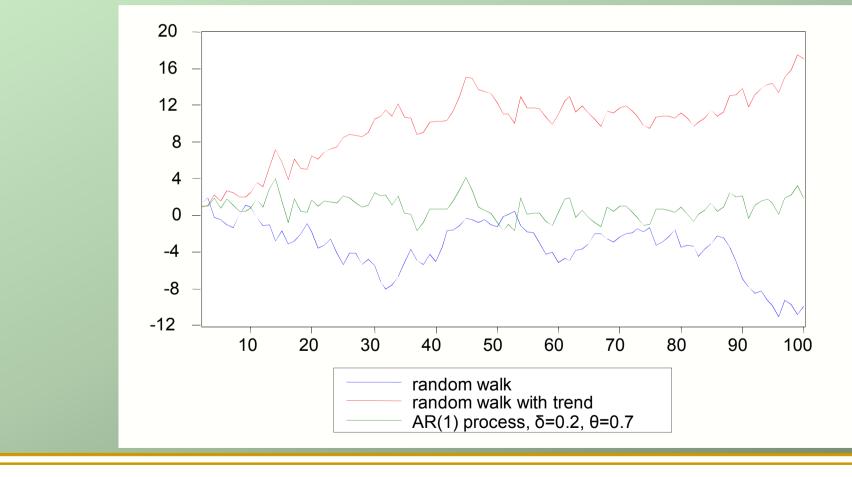
Stochastic trend:
$$Y_t = \delta + Y_{t-1} + \varepsilon_t$$
 or

 $\Delta Y_t = Y_t - Y_{t-1} = \delta + \varepsilon_t$, ε_t : white noise

- $\hfill\square$ describes an irregular or random fluctuation of the differences ΔY_t around the expected value δ
- AR(1) or AR(p) process with unit root
- "random walk with trend"

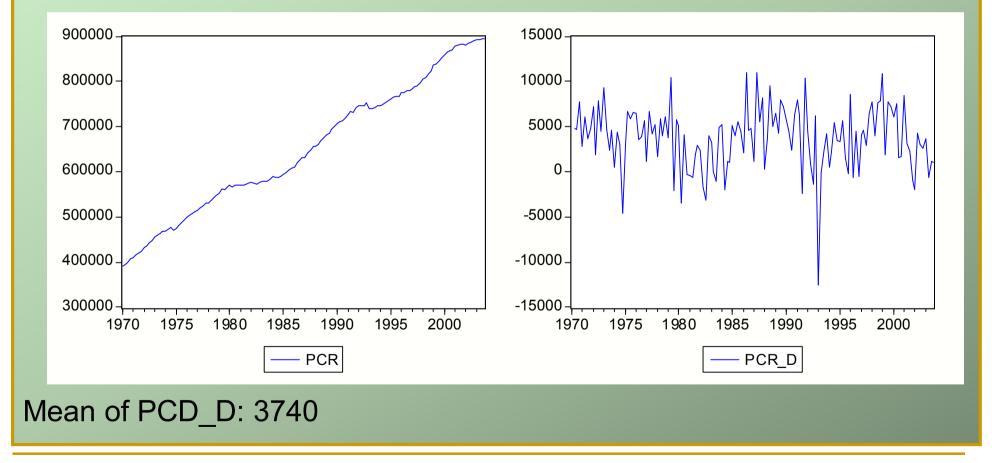
Trends: Random Walk and AR Process

Random walk: $Y_t = Y_{t-1} + \varepsilon_t$; random walk with trend: $Y_t = 0.1 + Y_{t-1} + \varepsilon_t$; AR(1) process: $Y_t = 0.2 + 0.7Y_{t-1} + \varepsilon_t$; ε_t simulated from N(0,1)



Example: Private Consumption

Private consumption, AWM database; level values (PCR) and first differences (PCR_D); random walk?



How to Model Trends?

Specification of a

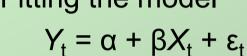
- deterministic trend, e.g., $Y_t = \alpha + \beta t + \varepsilon_t$: risk of spurious regression, wrong decisions
- stochastic trend: analysis of differences ΔY_t if a random walk, i.e., a unit root, is suspected

Spurious Regression: An Illustration

Independent random walks: $Y_t = Y_{t-1} + \varepsilon_{yt}$, $X_t = X_{t-1} + \varepsilon_{xt}$

-15 L

 ϵ_{yt} , ϵ_{xt} : independent white noises with variances $\sigma_y^2 = 2$, $\sigma_x^2 = 1$ Fitting the model

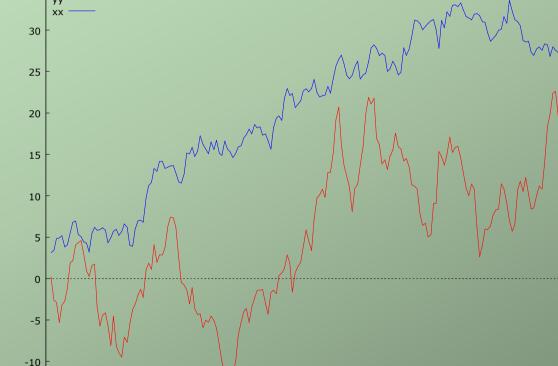


gives

$$\hat{Y}_{t} = -8.18 + 0.68X_{t}$$

-statistic for X. $t = 17.1$

$$p$$
-value = 1.2 E-40
 $\mathbb{R}^2 = 0.50$, $\mathbb{OW} = 0.12$



100

150

50

200

Models in Non-stationary Time Series

Given that $X_t \sim I(1)$, $Y_t \sim I(1)$ and the model is

 $Y_t = \alpha + \beta X_t + \varepsilon_t$

it follows in general that $\varepsilon_t \sim I(1)$, i.e., the error terms are nonstationary

Consequences for OLS estimation of α and β

- (Asymptotic) distributions of t- and F -statistics are not the t- and F distribution
- *t*-statistic, R² indicate explanatory potential
- Highly autocorrelated residuals, DW statistic converges for growing length of time series to zero

Nonsense or spurious regression (Granger & Newbold, 1974)

 Non-stationary time series are trended; non-stationarity causes an apparent relationship

Avoiding Spurious Regression

- Identification of non-stationarity: unit-root tests
- Models for non-stationary variables
 - Elimination of stochastic trends: specifying the model for differences
 - Inclusion of lagged variables may result in stationary error terms
 - Explained and explanatory variables may have a common stochastic trend, are cointegrated: equilibrium relation, error-correction models

Unit Root Tests

AR(1) process $Y_t = \delta + \theta Y_{t-1} + \varepsilon_t$ with white noise ε_t

- Dickey-Fuller or DF test (Dickey & Fuller, 1979) Test of H_0 : θ = 1 against H_1 : θ < 1
- KPSS test (Kwiatkowski, Phillips, Schmidt & Shin, 1992) Test of H₀: θ < 1 against H₁: θ = 1
- Augmented Dickey-Fuller or ADF test extension of DF test
- Various modifications like Phillips-Perron test, Dickey-Fuller GLS test, etc.

The Error-correction Model

ADL(1,1) model with $Y_t \sim I(1), X_t \sim I(1)$

 $Y_t = \delta + \Theta Y_{t-1} + \varphi_0 X_t + \varphi_1 X_{t-1} + \varepsilon_t$

Common trend implies an equilibrium relation, i.e.,

 $Y_{t-1} - \beta X_{t-1} \sim I(0)$

error-correction form of the ADL(1,1) model

$$\Delta Y_{t} = \varphi_{0} \Delta X_{t} - (1 - \theta) (Y_{t-1} - \alpha - \beta X_{t-1}) + \varepsilon_{t}$$

Error-correction model describes

- the short-run behaviour
- consistently with the long-run equilibrium $Y_t = \alpha + \beta X_t$

Cointegration

Non-stationary variables $X_t \sim I(1)$, $Y_t \sim I(1)$

 $Y_t = \alpha + \beta X_t + \varepsilon_t$

- X_t and Y_t are cointegrated: $\varepsilon_t \sim I(0)$
- X_t and Y_t are not cointegrated: $\varepsilon_t \sim I(1)$

Tests for cointegration:

- If β is known, unit root test based on differences $Y_t \beta X_t$
- Test procedures
 - \Box Unit root test (DF or ADF) based on residuals e_t
 - Cointegrating regression Durbin-Watson (CRDW) test: DW statistic
 - Johansen technique: extends the cointegration technique to the multivariate case

Vector Error-Correction Model

 Y_t : *k*-vector, each component *I*(1)

VAR(p) model for the *k*-vector Y_t

$$Y_{t} = \delta + \Theta_{1}Y_{t-1} + \dots + \Theta_{p}Y_{t-p} + \varepsilon_{t}$$

transformed into

$$\Delta Y_{t} = \delta + \Gamma_{1} \Delta Y_{t-1} + \dots + \Gamma_{p-1} \Delta Y_{t-p+1} + \Pi Y_{t-1} + \varepsilon_{t}$$

with $r{\Pi} = r$ and $\Pi = \gamma\beta'$ gives

$$\Delta Y_{t} = \delta + \Gamma_{1} \Delta Y_{t-1} + \dots + \Gamma_{p-1} \Delta Y_{t-p+1} + \gamma \beta' Y_{t-1} + \varepsilon_{t}$$
(B)

- *r* cointegrating relations $\beta' Y_{t-1}$
- Adaptation parameters γ measure the portion or speed of adaptation of Y_t in compensation of the equilibrium error $Z_{t-1} = \beta' Y_{t-1}$
- Equation (B) is called the vector error-correction (VEC) form of the VAR(p) model

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Example

To be explained whether a household owns a car: explanatory power have

- income
- household size
- etc.

Regression for describing car-ownership is not suitable!

- Owning a car has two manifestations: yes/no
- Indicator for owning a car is a binary variable

Models are needed that allow describing a binary dependent variable or a, more generally, limited dependent variable

Cases of Limited Dependent Variable

- Typical situations: functions of explanatory variables are used to describe or explain
- Dichotomous dependent variable, e.g., ownership of a car (yes/no), employment status (employed/unemployed), etc.
- Ordered response, e.g., qualitative assessment (good/average/bad), working status (full-time/part-time/not working), etc.
- Multinomial response, e.g., trading destinations (Europe/Asia/Africa), transportation means (train/bus/car), etc.
- Count data, e.g., number of orders a company receives in a week, number of patents granted to a company in a year
- Censored data, e.g., expenditures for durable goods, duration of study with drop outs

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Panel Data

Population of interest: individuals, households, companies, countries

- Types of observations
- Cross-sectional data: Observations of all units of a population, or of a (representative) subset, at one specific point in time; e.g., wages in 2015
- Time series data: Series of observations on units of the population over a period of time; e.g., wages of a worker in 2009 through 2015
- Panel data: Repeated observations of (the same) population units collected over a number of periods; data set with both a cross-sectional and a time series aspect; multi-dimensional data
- Cross-sectional and time series data are one-dimensional, special cases of panel data

Pooling independent cross-sections: (only) similar to panel data

Panel Data: Three Types

Typically data at micro-economic level (individuals, households, firms), but also at macro-economic level (e.g., countries) Notation:

- *N*: Number of cross-sectional units
- *T*: Number of time periods

Types of panel data:

- Large T, small N: "long and narrow"
- Small *T*, large *N*: "short and wide"
- Large *T*, large *N*: "long and wide"

Some Examples

Verbeek's data set "males": Wages and related variables

- short and wide panel (N = 545, T = 8)
- rich in information (~40 variables)

Grunfeld investment data: Investments in plant and equipment by

- N = 10 firms
- for each of T = 20 yearly observations for 1935-1954
- Penn World Table: Purchasing power parity and national income accounts for
- N = 189 countries/territories
- for some or all of the years 1950-2009 ($T \le 60$)

Example: Individual Wages

Verbeek's data set "males"

- Sample of
 - 545 full-time working males, end of schooling in 1980
 - from each person: yearly data collection from 1980 till 1987
- Variables
 - wage: log of hourly wage (in USD)
 - school: years of schooling
 - exper: age 6 school
 - dummies for union membership, married, black, Hispanic, public sector
 - others

Use of Panel Data

Econometric models for describing the behaviour of cross-sectional units over time

Panel data models

- Allow controlling individual differences, comparing behaviour, analysing dynamic adjustment, measuring effects of policy changes
- More realistic models than cross-sectional and time-series models
- Allow more detailed or sophisticated research questions

Methodological implications

- Dependence of sample units in time-dimension
- Some variables might be time-constant (e.g., variable school in "males", population size in the Penn World Table dataset)
- Missing values

Examples for Fixed- and Random-effects

Grunfeld investment data: Investment model

 $I_{it} = \alpha_i + \beta_{i1}F_{it} + \beta_{i2}C_{it} + u_{it}$

with *F*_{it}: market value, *C*_{it}: value of stock of plant and equipment, both of firm *i* at the end of year *t*-1

• N = 10 firms, T = 20 yearly observations

• Fixed effects α_i allow for firm-specific, time-constant factors Wage equation

 $wage_{it} = \beta_1 + \beta_2 exper_{it} + \beta_3 exper_{it} + \beta_4 school_{it} + \beta_5 union_{it}$

+ $\beta_6 mar_{it}$ + $\beta_7 black_{it}$ + $\beta_8 rural_{it}$ + α_i + u_{it}

with composite error $\varepsilon_{it} = \alpha_i + u_{it}$

 \circ α_i : unit-specific parameter for each of 545 units

- Time-constant factors α_i : stochastic variables with identical distribution
- Regressors are uncorrelated with u_{it}

Models for Panel Data

Model for *y*, based on panel data from *N* cross-sectional units and *T* periods

$$y_{it} = \beta_0 + x_{it}'\beta_1 + \varepsilon_{it}$$

- *i* = 1, ..., *N*: sample unit
- t = 1, ..., T: time period of sample

 x_{it} and β_1 : *K*-vectors

- $β_0$ and $β_1$: represent intercept and K regression coefficients; are assumed to be identical for all units and all time periods
- ε_{it} : represents unobserved factors that may affect y_{it}
 - Assumption that ε_{it} are uncorrelated over time not realistic; refer to the same unit or individual
 - Standard errors of OLS estimates misleading, OLS estimation not efficient (does not exploit dependence structure over time)

Fixed Effects Model

The general model

 $y_{it} = \beta_0 + x_{it}'\beta_1 + \varepsilon_{it}$

Specification for the error terms: two components

 $\varepsilon_{\rm it} = \alpha_{\rm i} + u_{\rm it}$

- \Box α_i fixed, unit-specific, time-constant factors, also called unobserved (individual) heterogeneity; may be correlated with x_{it}
- □ $u_{it} \sim IID(0, \sigma_u^2)$; homoskedastic, uncorrelated over time; represents unobserved factors that change over time, also called idiosyncratic or time-varying error
- ε_{it} : also called composite error
- Fixed effects (FE) model

 $y_{it} = \sum_{j} \alpha_{i} d_{ij} + x_{it}'\beta_{1} + u_{it}$

 d_{ij} : dummy variable for unit *i*: $d_{ij} = 1$ if i = j, otherwise $d_{ij} = 0$

• Overall intercept β_0 omitted; unit-specific intercepts α_i

Fixed Effects Estimator

Within transformation": transforms y_{it} into time-demeaned \ddot{y}_{it} by subtracting the average $\bar{y_i} = (\Sigma_t y_{it})/T$: $\ddot{y}_{it} = y_{it} - \bar{y_i}$; analogously \ddot{x}_{it} and \ddot{u}_{it} , for all *i* and *t*

 $b_{\mathsf{FE}} = (\Sigma_i \Sigma_t \, \ddot{X}_{it} \, \ddot{X}_{it}')^{-1} \, \Sigma_i \Sigma_t \, \ddot{X}_{it} \, \ddot{y}_{it}$

- Unbiased if all x_{it} are independent of all u_{it}
- Consistent (for $N \to \infty$) if x_{it} are strictly exogenous, i.e., $E\{x_{it} u_{is}\} = 0$ for all s, t
- Asymptotically normally distributed
- Covariance matrix

 $V\{b_{FE}\} = \sigma_u^2 (\Sigma_i \Sigma_t \ddot{X}_{it} \ddot{X}_{it})^{-1}$

Random Effects Model

Starting point is again the model

 $y_{it} = \beta_0 + x_{it} \beta_1 + \varepsilon_{it}$

with composite error $\varepsilon_{it} = \alpha_i + u_{it}$

Specification for the error terms:

- □ $u_{it} \sim IID(0, \sigma_u^2)$; homoskedastic, uncorrelated over time
- α_i ~ IID(0, σ_a^2); represents all unit-specific, time-constant factors; correlation of error terms over time only via the α_i
- α_i and u_{it} are assumed to be mutually independent and independent of x_{is} for all *j* and *s*
- Random effects (RE) model

 $y_{it} = \beta_0 + x_{it}'\beta_1 + \alpha_i + u_{it}$

- Unbiased and consistent (N $\rightarrow \infty$) estimation of β_0 and β_1
- Efficient estimation of β_0 and β_1 : takes error covariance structure into account; GLS estimation

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Econometrics II

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