

Cauchyho věta

$$|(x, y)|^2 \leq \|x\|^2 \cdot \|y\|^2$$

odkud $\cos(x, y) = \frac{(x, y)}{\|x\| \cdot \|y\|}$

$$(x_1 y_1 + x_2 y_2 + \dots + x_n y_n)^2 \leq (x_1^2 + x_2^2 + \dots + x_n^2)(y_1^2 + y_2^2 + \dots + y_n^2)$$

$$y = a + bX$$

$$M_y(t) = E e^{t \cdot y} = E \cdot e^{t(a + b \cdot X)} =$$

$$= E e^{t \cdot a} \cdot e^{t \cdot b \cdot X} = e^{t \cdot a} \cdot \underbrace{E e^{t \cdot b \cdot X}}_n$$

$$= e^{t \cdot a} M_x(t \cdot b)$$

$$\Rightarrow Z \sim N(0, 1) \\ Y \sim a + bZ$$

$$M_Z(t) = e^{-\frac{t^2}{2}}$$

$$M_Y(t) = e^{at} \cdot e^{-\frac{b^2 t^2}{2}}$$

$$Z \sim N(0,1)$$

$$f_Z(z) = \frac{1}{\sqrt{2\pi}} \cdot e^{-\frac{z^2}{2}}$$

$$M_Z(t) = E[e^{t \cdot Z}] = \int_{-\infty}^{\infty} e^{t \cdot z} \cdot f_Z(z) dz$$

$$= \int_{-\infty}^{\infty} e^{t \cdot z} \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} dz$$

$$\Rightarrow M_Z(t) = e^{\frac{t^2}{2}}$$

$E[Z] = 0$... $\text{var}[Z] = 1$ a $\text{var}[Z] = 1$ do exp. vidy

$$e^x = \sum_{k=0}^{\infty} \frac{x^k}{k!} = 1 + \frac{x}{1!} + \frac{x^2}{2!} + \frac{x^3}{3!} + \dots$$

$$\Rightarrow e^{2t^2} = 1 + \frac{t^2}{2!} + \frac{4t^4}{4!} + \dots + \frac{t^{2k}}{k!} + \dots$$

$$\Rightarrow \mu_1^- = 0 \quad (\text{coef. } \sim t^{-1}) \Rightarrow E Z = 0$$

$$\begin{aligned} \mu_2^- &= 1 \quad (\text{coef. } \sim t^2) \Rightarrow DZ = E(Z^2) - \\ &\quad - E(Z)^2 = \\ &\quad - \mu_2^- - (\mu_1^-)^2 = 1 \end{aligned}$$

$$\begin{aligned} \Gamma\left(\frac{1}{2}, \frac{1}{2}\right) &= \sqrt{\frac{1}{2}} \cdot \frac{1}{\sqrt{x}} \cdot x^{-\frac{1}{2}} e^{-\frac{1}{2}x} \\ &= \frac{1}{\sqrt{2x}} \cdot \frac{1}{\sqrt{x}} e^{-\frac{1}{2}x} \end{aligned}$$

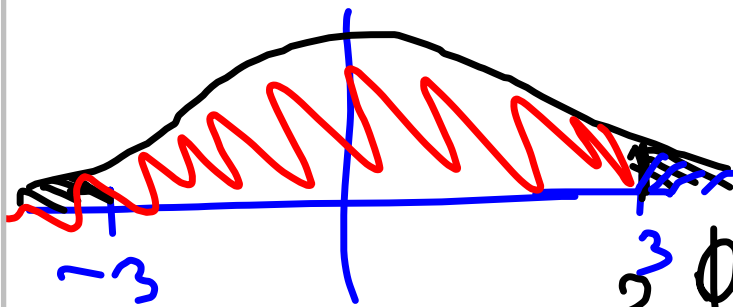
$$\varepsilon := k \cdot \frac{\sqrt{\text{DX}}}{\sigma} \dots \text{sm. odległość}$$

$$EX =: \mu$$

$$P(|X - \mu| \geq k \cdot \sigma) \leq \frac{\sigma^2}{k^2 \sigma^2} = \frac{1}{k^2}$$

$$\underline{k=3}: P(|X - \mu| \geq 3\sigma) \leq \frac{1}{9}$$

$$\begin{aligned} X \sim N(0,1): P(|X - \mu| \geq 3\sigma) &= \\ P(|X| \geq 3) &= 2\Phi(-3) = 2(1 - \Phi(3)) = \\ &= 2(0,0044) = \\ &= \underline{\underline{0,0088}} \end{aligned}$$



$$2\Phi(-3) = 2(1 - \Phi(3))$$

Věta (Bernoulliiova)

Pro náhodnou veličinu s binomickým rozdělením $Y_n \sim \text{Bi}(n, p)$ a pro libovolné $\epsilon > 0$ platí

$$P\left(\left|\frac{Y_n}{n} - p\right| > \epsilon\right) \leq \frac{p(1-p)}{n\epsilon^2}.$$

plyne
z Ceb. nerovnosti

$$\lim_{n \rightarrow \infty} P\left(\left|\frac{Y_n}{n} - p\right| > \epsilon\right) \leq \lim_{n \rightarrow \infty} \frac{p(1-p)}{n\epsilon^2} = 0$$
$$\lim_{n \rightarrow \infty} P\left(\left|\frac{Y_n}{n} - p\right| \leq \epsilon\right) = 1$$

$$Y_n = \sum_{i=1}^n A_i, \text{ kde } A_i \sim B(1, p) = A(p)$$

$$EA = p$$

$$DA = p(1-p)$$

$$\lim_{n \rightarrow \infty} P \left(\left| \frac{1}{n} \cdot \sum_{i=1}^n A_i - p \right| < \varepsilon \right) = 1$$

$$\stackrel{||}{=} \lim_{n \rightarrow \infty} P \left(\left| \frac{\sum_{i=1}^n A_i - np}{n} \right| < \varepsilon \right) = 1$$

$$\stackrel{||}{=} P \left(\left| \frac{Y_n}{n} - p \right| < \varepsilon \right) \Rightarrow 1$$

Bernoulliova

Př: $X_n \dots$ # studenti, využívající nebo nevyužívající

$$X_n \sim \text{Bi}(n, p)$$

$p \dots$ nezávadná $n = 600$

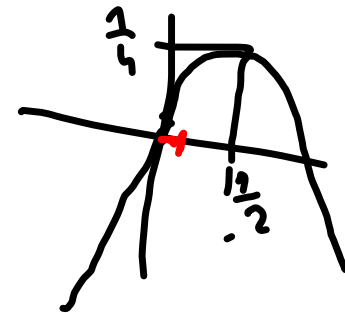
$$p \leq 0,02 =: p_0 \quad \epsilon = 0,01 = \frac{1}{100}$$

Bernoulli

$$P\left(\left|\frac{X_n}{n} - p\right| > \epsilon\right) \leq \frac{p(1-p)}{n \cdot \epsilon^2}$$

$$p(1-p) \leq p_0(1-p_0) = 0,02 \cdot 0,98$$

$$P\left(\left|\frac{X_n}{n} - p\right| > \frac{1}{100}\right) \leq \frac{p(1-p)}{n \cdot 10^{-4}}$$



$f(x) = x(1-x)$
roste pro $x \in (0, 1/2)$

$$\leq \frac{2 \cdot 10^{-2} \cdot 98 \cdot 10^{-2}}{600 \cdot 10^4} =$$

$$\approx \frac{2 \cdot 10^{-2}}{6 \cdot 10^{-2}} = \frac{1}{3}$$

$$P\left(\left|p - \frac{X}{n}\right| > \frac{1}{100}\right) \leq \frac{1}{3}$$

$$Y_1, Y_2, \dots, Y_n, \dots$$

$$X_1 = Y_1, X_2 = Y_1 + Y_2, \dots, X_n = \sum_{i=1}^n Y_i$$

$$EX_n = n \cdot \mu$$

$$DX_n = n \cdot \sigma^2$$

$$D(x+y) = Dx + Dy$$

normaliz: $\frac{X_n - n\mu}{\sqrt{n \cdot \sigma^2}} = \frac{1}{\sqrt{n}} \cdot \frac{X_n - n\mu}{\sigma} =$

$$= \frac{1}{\sqrt{n}} \cdot \frac{\sum_{i=1}^n Y_i - n\mu}{\sigma} = \frac{1}{\sqrt{n}} \sum_{i=1}^n \frac{Y_i - \mu}{\sigma}$$

Centralni limit
 $N(0,1)$