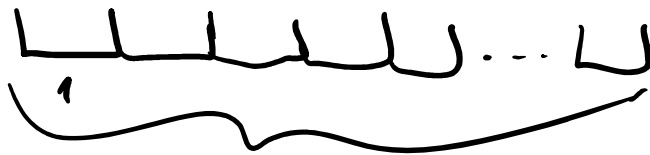


Basic methods: Moments and Deviations

Occupancy problem (bins and balls)

Coupon collector problem

$m$  - balls  
 $n$  - bins



expected

[bin number 1] - What is the  $V$  number of balls you need to place before one of them lands in bin 1.

Geometric distribution  $\lambda$

1  
2  
⋮  
n

Probability 1st ball falls into 1st bin =  $\frac{1}{n}$

Probability 2nd ball is the first one to land into bin 1

$$\frac{1}{n} \cdot \frac{n-1}{n}$$

Probability that  $n$ -th ball is the first one to land into bin 1

n } Probability that n-th ball is the first one to land into bin 1

$$E(X) = \frac{1}{p} \quad \text{if the probability success is } p.$$

$$\begin{aligned} E(X) &= \sum_{i=1}^{\infty} i \cdot P(X=i) \\ &= \sum_{i=1}^{\infty} i \cdot (1-p)^{i-1} \cdot p \\ &= p \cdot \sum_{i=1}^{\infty} i \cdot (1-p)^{i-1} \quad / \cdot (1-p) \end{aligned}$$

$$(1-p) \cdot E(X) = p \cdot \sum_{i=1}^{\infty} i \cdot (1-p)^i$$

$$\begin{aligned} (1 - (1-p))E(X) &= p \cdot \sum_{i=1}^{\infty} i \cdot (1-p)^{i-1} - p \cdot \sum_{i=1}^{\infty} i \cdot (1-p)^i \\ &= p \left[ \underbrace{1 - (1-p)}_{i=1} + \underbrace{2 \cdot (1-p) - 2 \cdot (1-p)^2}_{i=2} \right. \\ &\quad \left. + \underbrace{3 \cdot (1-p)^2 - 3 \cdot (1-p)^3}_{i=3} \right] \end{aligned}$$

$$= p \left[ 1 + (1-p) + (1-p)^2 + (1-p)^3 + \dots + \right]$$

$$\underbrace{\quad}_{=1}$$

$$\sum_{i=1}^{\infty} v^i = \frac{1}{1-v}$$

$$= p \cdot \frac{1}{1 - (1-p)}$$

$$p \cdot E(x) = p \cdot \frac{1}{p}$$

$$E(x) = \frac{1}{p}$$

$$\sum_{i=0}^{\infty} v^i = \frac{1}{1-v}$$

$$|v| < 1$$

$m$  balls and  $n$  bins

What is the expected number of empty bins.

$X_i = 1$  if  $i^{\text{th}}$  bin is empty

$$P_r(X_i = 1) = \left(\frac{n-1}{n}\right)^m$$

$$E(X_i) = \left(\frac{n-1}{n}\right)^m$$

$$X = \sum_i X_i$$

$$E(X) = \sum_{i=1}^n E(X_i) = n \cdot \left(\frac{n-1}{n}\right)^m$$

DRUNKEN SAILOR PROBLEM ↑

(coupon collector problem)

I have 0 coupons  $\rightarrow$  Probability that the next coupon I get is new?

$$Pr = 1$$

I have 1 coupon  $\rightarrow$   $Pr(\text{New}) = \frac{n-1}{n}$   
(2nd)

2 coupons  $\rightarrow$   $Pr(\text{new}) = \frac{n-2}{n}$   
(3rd)

$i^{\text{th}}$  coupon  $\rightarrow$   $Pr(\text{new}) = \frac{n-i+1}{n}$   
( $(i+1)^{\text{st}}$ )

these are called

"Even" in the slides

$X_i$   $\rightarrow$  the average number of coupons to get  $i+1^{\text{st}}$  coupon

these are geometric distributions with  $p_i = \left(\frac{n-i+1}{n}\right)$

$$X = \sum_i X_i$$

$$E(X) = E\left(\sum_i X_i\right) = \sum_i E(X_i)$$

$$= \sum_{i=1}^n \frac{1}{p_i} = \sum_{i=1}^n \frac{n}{n-i} = n \cdot \sum_{i=1}^n \frac{1}{n-i+1}$$

$$= n \cdot \sum_{i=1}^n \frac{1}{i} = n \cdot H_n \approx n \cdot \log n$$

$\Sigma_i(k) =$  event that  $j^{\text{th}}$  bin has  $k$  or more balls  
.. "

$\Sigma_j(k)$  = event that  $j$ th bin has  $k$  or more balls  
 $n$ -balls and  $n$ -bins

$$\begin{aligned} \Pr(\text{jth bin has exactly } i \text{ balls}) &= \binom{n}{i} \left(\frac{1}{n}\right)^i \left(1 - \frac{1}{n}\right)^{n-i} \\ &\leq \binom{n}{i} \left(\frac{1}{n}\right)^i \\ &\leq \left(\frac{ne}{i}\right)^i \left(\frac{1}{n}\right)^i \quad \left(\binom{n}{i} \leq \left(\frac{ne}{i}\right)^i\right) \\ &= \left(\frac{e}{i}\right)^i \end{aligned}$$

$$\Pr \Sigma_j(k) = \sum_{i=k}^n \binom{n}{i} \left(\frac{1}{n}\right)^i \left(\frac{n-1}{n}\right)^{n-i}$$

$$\leq \sum_{i=k}^n \left(\frac{e}{i}\right)^i = \left(\frac{e}{k}\right)^k + \left(\frac{e}{k+1}\right)^{k+1} + \left(\frac{e}{k+2}\right)^{k+2} + \dots$$

$$\leq \left(\frac{e}{k}\right)^k + \left(\frac{e}{k}\right)^{k+1} + \left(\frac{e}{k}\right)^{k+2} + \dots$$

$$= \left(\frac{e}{k}\right)^k \left(1 + \left(\frac{e}{k}\right) + \left(\frac{e}{k}\right)^2 + \dots + \left(\frac{e}{k}\right)^n\right)$$

$$\leq \left(\frac{e}{k}\right)^k \sum_{i=0}^{\infty} \left(\frac{e}{k}\right)^i$$

if  $\left(\frac{e}{k}\right) < 1$

$$\leq \left(\frac{e}{k}\right)^k \cdot \frac{1}{1 - \frac{e}{k}}$$

for  $k = \left\lceil \frac{e \ln n}{\ln \ln n} \right\rceil$

$$\leq \frac{1}{n^2}$$

I couldn't reach this =)

for  $\epsilon = \frac{1}{\ln \ln n}$   $\left[ \frac{1}{n^2} \right]$  (condition must hold -)

$$\Pr \left[ \bigcup_{i=1}^n \epsilon_i(k) \right] \leq \sum_{i=1}^n \Pr \epsilon_i(k) = \frac{1}{n}$$