

M7777 Applied Functional Data Analysis

12. Sparse FDA

Jan Koláček (kolacek@math.muni.cz)

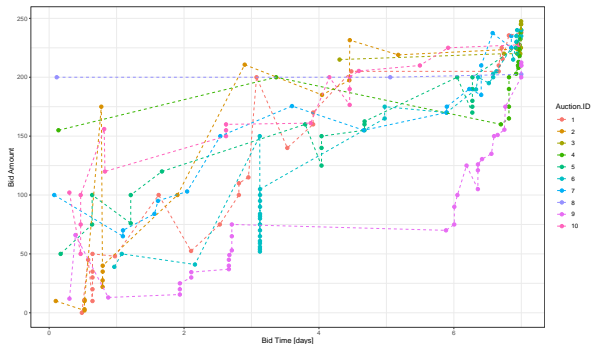
Dept. of Mathematics and Statistics, Faculty of Science, Masaryk University, Brno



Ebay Auctions

Jank and Shmueli, 2007

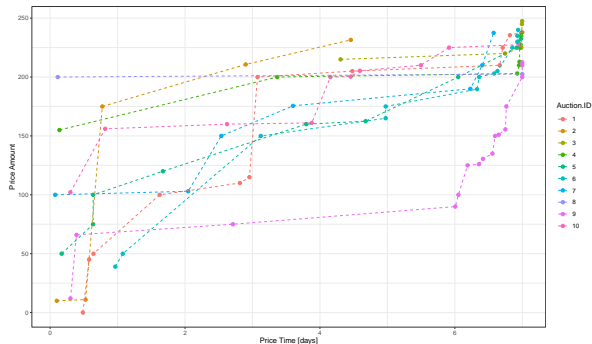
- 7-Day auctions for new Palm M515 PDAs
- 149 Auctions, collected May-June 2003



Sample of 10 auctions – bid histories.

Auction Price

- Only increases if bid is greater than current price



Sample of 10 auctions – price histories.

We will consider a model

$$Y_{ij} = \underbrace{\mu(t_{ij}) + \varepsilon_i(t_{ij})}_{X_i(t_{ij})} + \delta_{ij},$$

for $1 \leq i \leq n$, $1 \leq j \leq n_i$, with assumptions

$\mu(t)$... the mean function (required to be smooth)

$\varepsilon_i(t)$... subject specific error functions, induce correlation between observations on the same subject, let's denote $c(s, t) = \text{Cov}(X(s), X(t)) = \text{Cov}(\varepsilon(s), \varepsilon(t))$

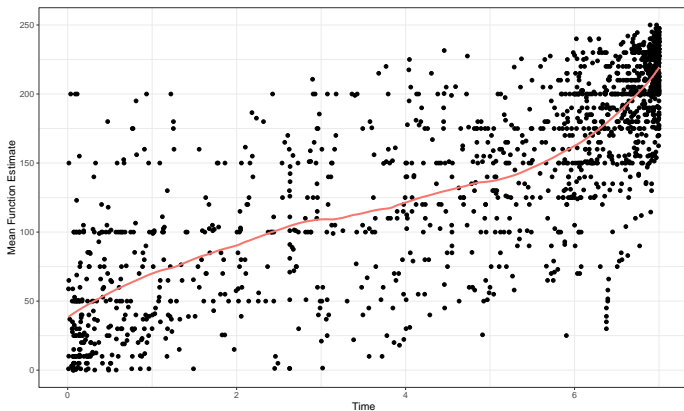
δ_{ij} ... errors explaining measurement noise, iid across both i and j , let's denote $\text{Var}(\delta_{ij}) = \sigma^2(t_{ij})$.

It means, that we observe a process $Y(t)$ in n samples $X_i(t)$, the i -th sample is observed in times t_1, \dots, t_{n_i} with setting

$$\text{Cov}(Y(s), Y(t)) = c(s, t) + \sigma^2(s)I_{s=t}.$$

The Main Idea

- 1 Let us consider all measurements Y_{ij} , $1 \leq i \leq n$, $1 \leq j \leq n_i$
- 2 Get an estimate $\hat{\mu}(t)$ of the mean function $\mu(t)$ (nonparametric, e.g. local linear kernel smoother, spline smoothing etc.)



- ③ Let us consider a set of time points pairs

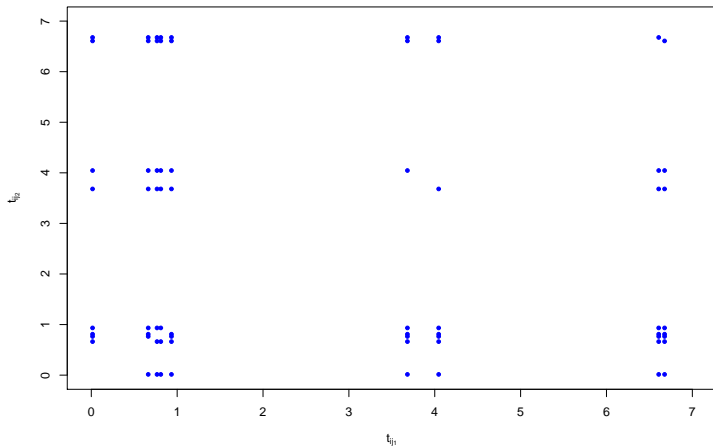
$$\mathbf{T} = \{(t_{ij_1}, t_{ij_2}) : 1 \leq i \leq n, 1 \leq j_1 \leq n_i, 1 \leq j_2 \leq n_i, j_1 \neq j_2\}$$

with its values

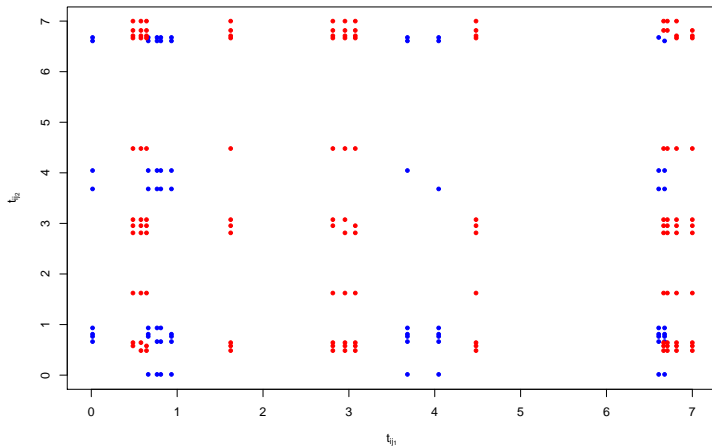
$$Z(t_{ij_1}, t_{ij_2}) = (Y_{ij_1} - \hat{\mu}(t_{ij_1}))(Y_{ij_2} - \hat{\mu}(t_{ij_2})), (t_{ij_1}, t_{ij_2}) \in \mathbf{T}$$

and get the covariance surface estimate $\hat{c}(s, t)$ (bivariate local linear etc.).

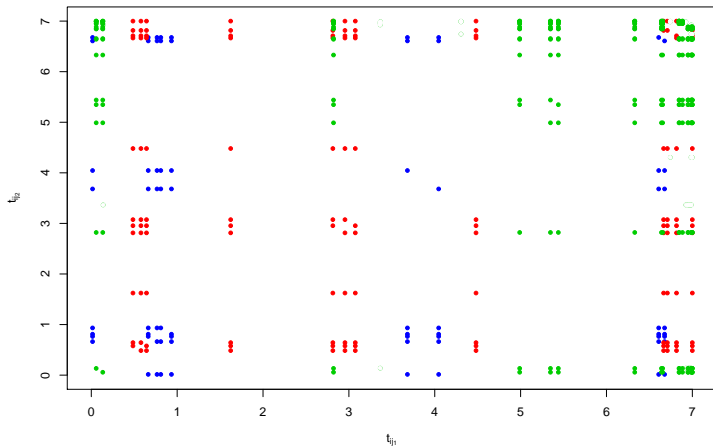
Samples: 1



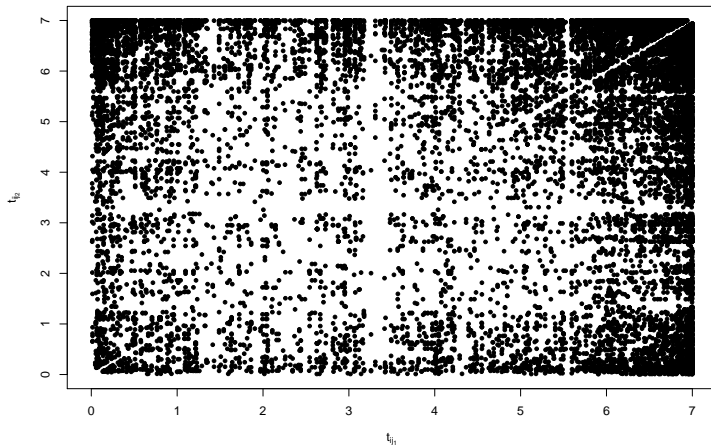
Samples: 2



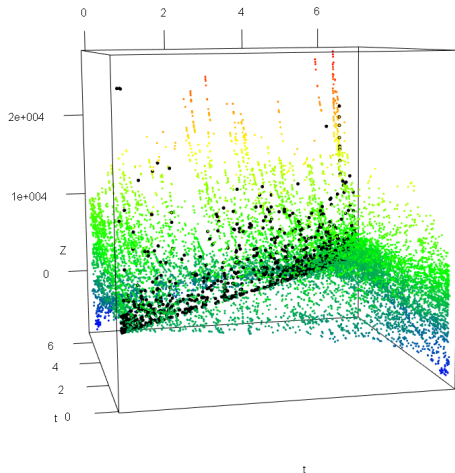
Samples: 3



Samples: all

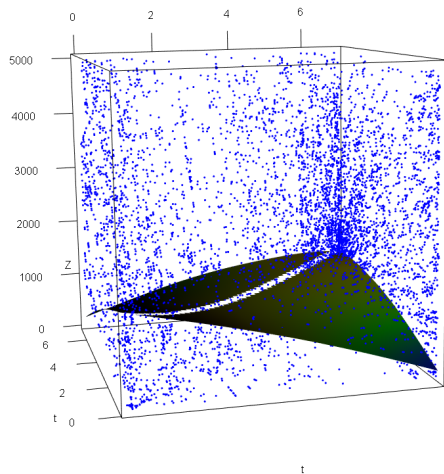


Auction Price



Raw Covariance Plot, black – diagonal terms

Auction Price



Covariance Estimate $\hat{c}(s, t)$, diagonal excluded

- 4 Take diagonal terms only

$$\mathbf{T}_{diag} = \{(t_{ij}, t_{ij}) : 1 \leq i \leq n, 1 \leq j \leq n_i\} \text{ and its } Z(t_{ij}, t_{ij})$$

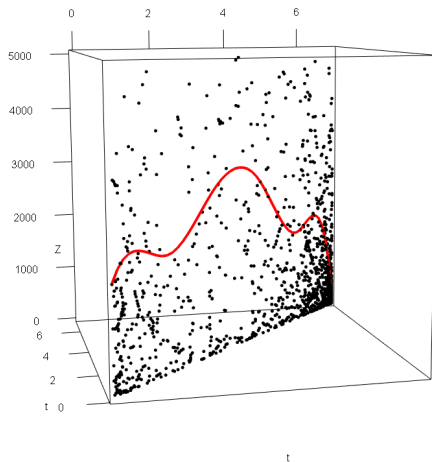
and by a univariate smoother get $\tilde{c}(t, t)$. Thus, an estimate of $\sigma^2(t)$

$$\hat{\sigma}^2(t) = \tilde{c}(t, t) - \hat{c}(t, t).$$

- 5 The estimate of $\text{Cov}(Y(s), Y(t))$ takes the form

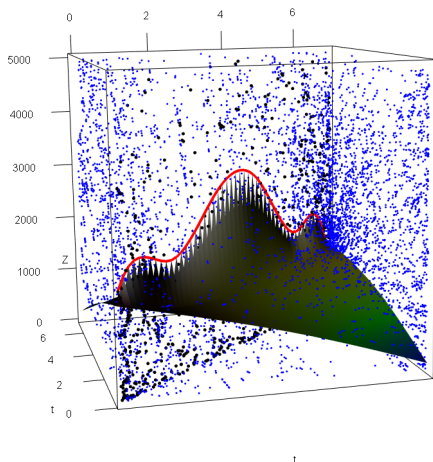
$$\hat{\sigma}(s, t) = \hat{c}(s, t) + \hat{\sigma}^2(t)$$

Auction Price



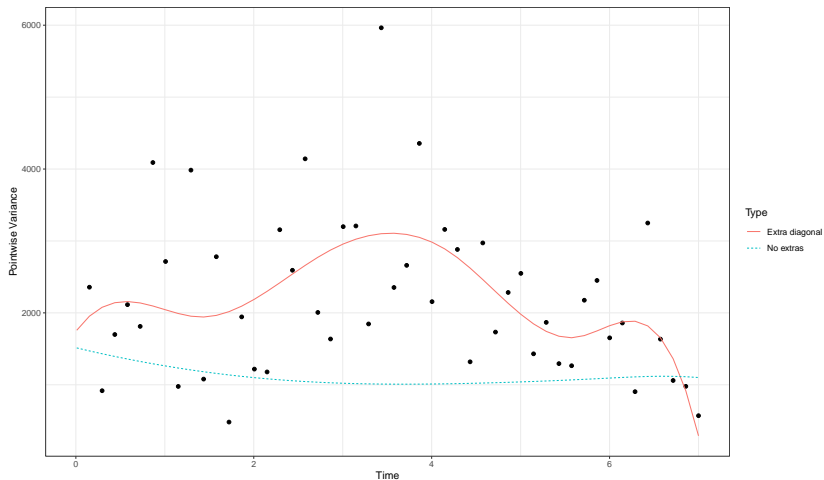
Variance Estimate $\hat{\sigma}^2$

Auction Price



$$\text{Covariance Estimate } \hat{\sigma}(s, t) = \hat{c}(s, t) + \hat{\sigma}^2(t)$$

Auction Price



Comparison of Variance Estimates

- ⑥ Let's consider the estimate of $\hat{\sigma}(s, t)$ and its Karhunen – Loève decomposition for functions

$$\hat{\sigma}(s, t) = \sum_{j=1}^{\infty} \lambda_j \xi_j(s) \xi_j(t) \quad \Rightarrow \quad \text{obtain } \hat{\xi}_j(t), \hat{\lambda}_j, j = 1, \dots, K.$$

- ⑦ Estimate principal scores $c_{ij} = \int \xi_j(t) [Y_i(t) - \mu(t)] dt$ through the conditional expectation

$$\hat{c}_{ij} = E[c_{ij} | \mathbf{Y}_i] = \hat{\lambda}_j \hat{\boldsymbol{\xi}}_j^T \hat{\boldsymbol{\Sigma}}_i^{-1} (\mathbf{Y}_i - \hat{\boldsymbol{\mu}}_i)$$

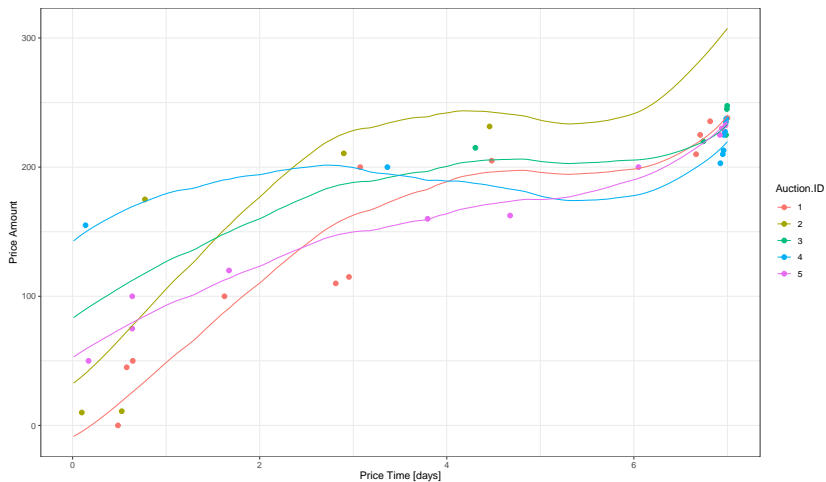
Yao et al. (2005)

- ⑧ Finally, reconstruct the whole curves

$$\hat{Y}_i(t) = \hat{\mu}(t) + \sum_{j=1}^K \hat{c}_{ij} \hat{\xi}_j(t).$$

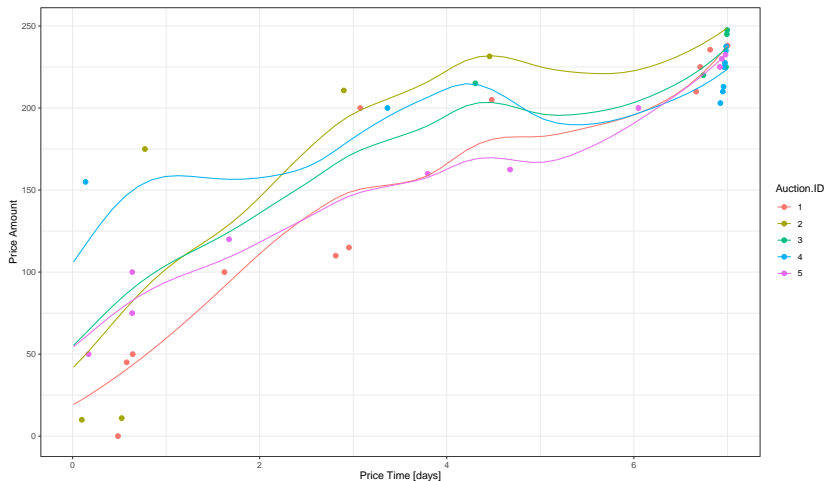
Filled Data

Auction Price – proposed method



Auction Prices Estimates

Auction Price – FDAPACE



Auction Prices Estimates

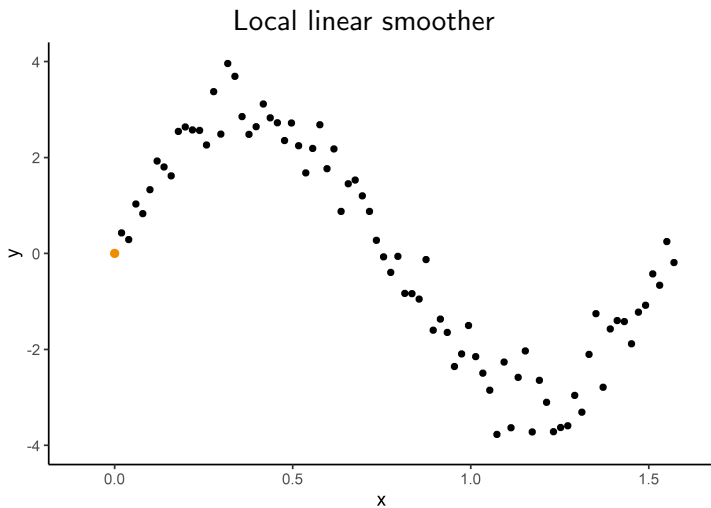
Mean function estimate

Local linear smoother with global bandwidth

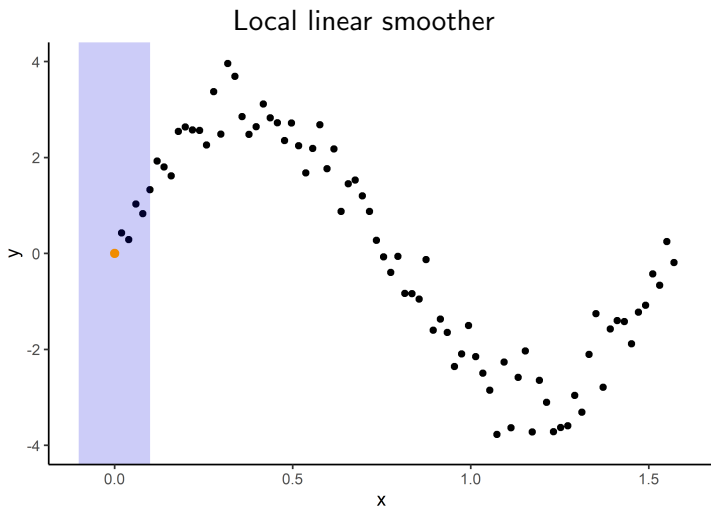
$$\sum_{i=1}^n \sum_{j=1}^{N_i} \left[K \left(\frac{T_{ij} - t}{h} \right) Y_{ij} - \beta_0 - \beta_1(t - T_{ij}) \right]^2 \rightarrow \min$$

- $K(x)$... **kernel** function (a symmetric density)
- h ... **global** bandwidth
- $\hat{\mu}(t) = \hat{\beta}_0(t)$

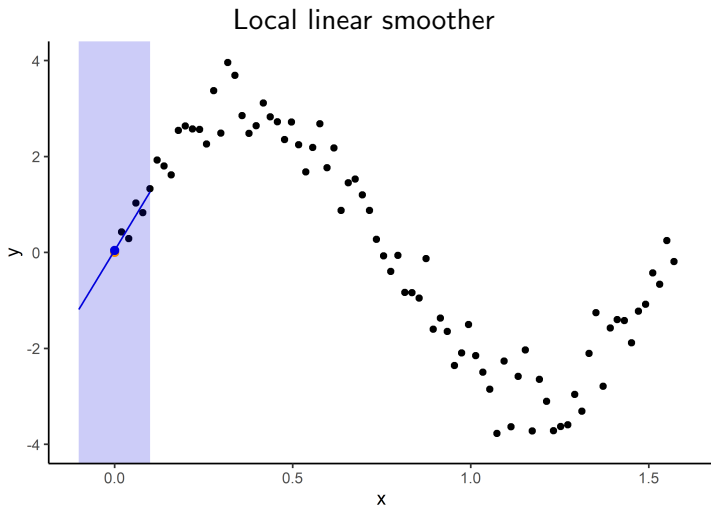
Kernel Smoothing



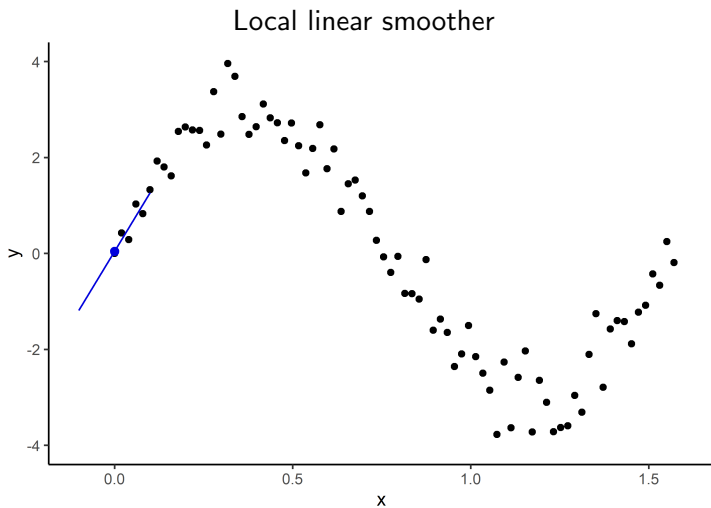
Kernel Smoothing



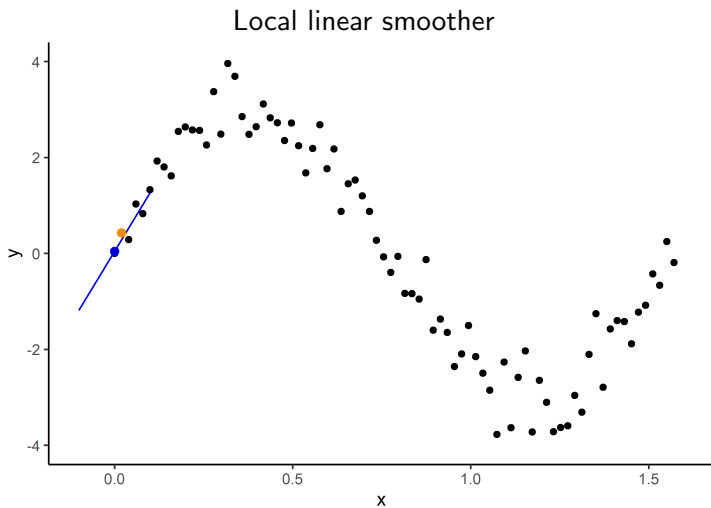
Kernel Smoothing



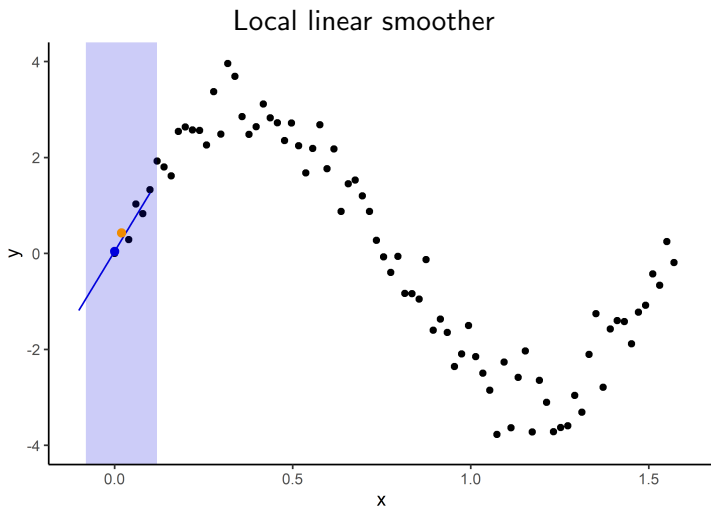
Kernel Smoothing



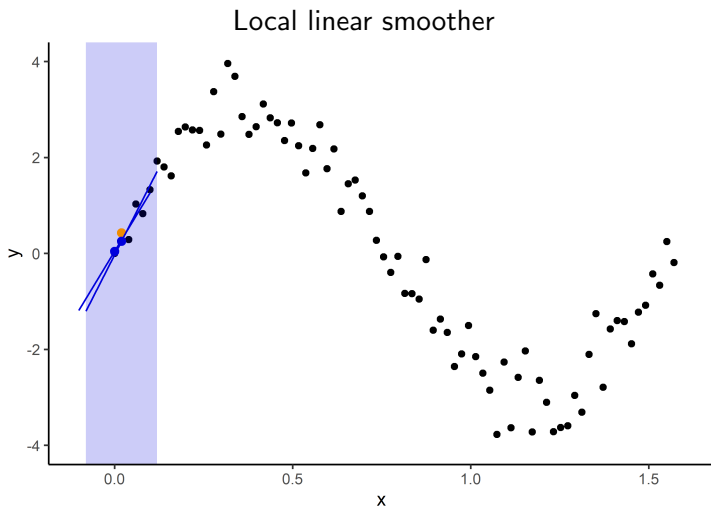
Kernel Smoothing



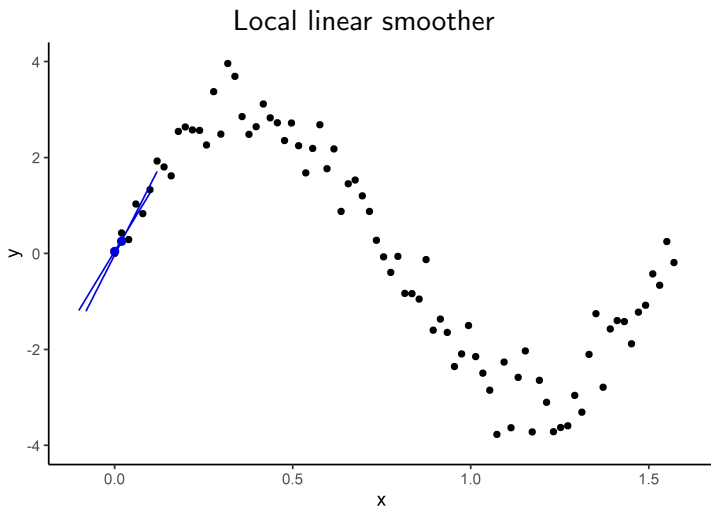
Kernel Smoothing



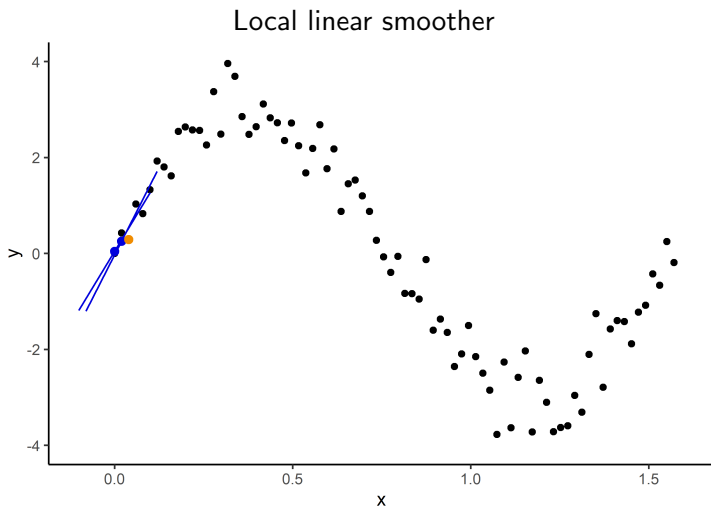
Kernel Smoothing



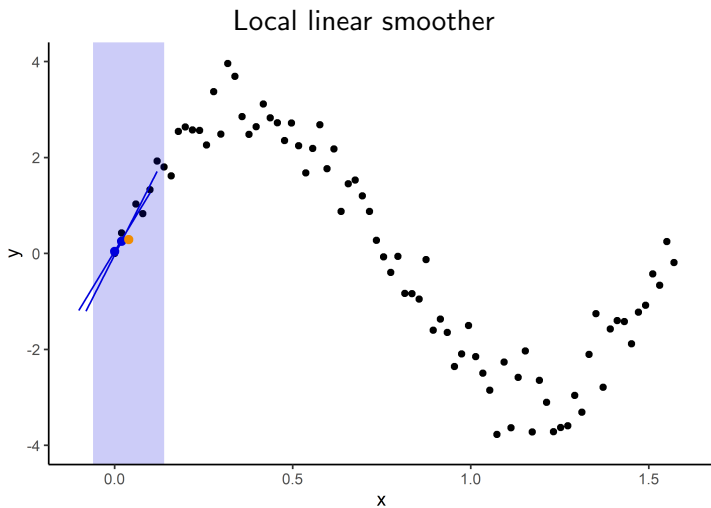
Kernel Smoothing



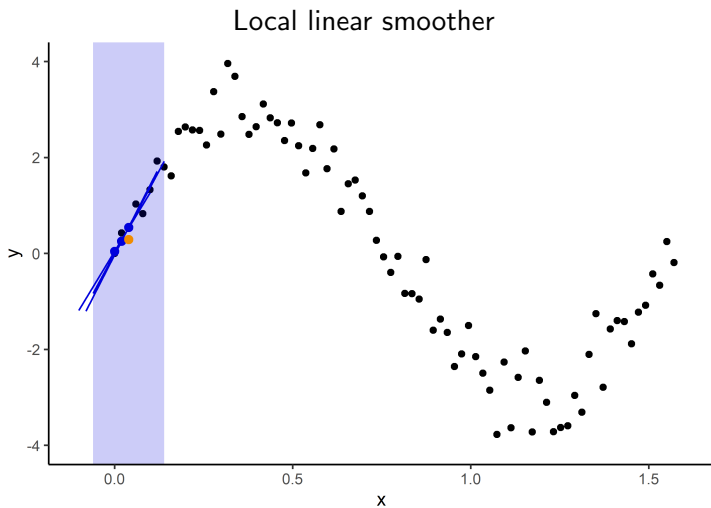
Kernel Smoothing



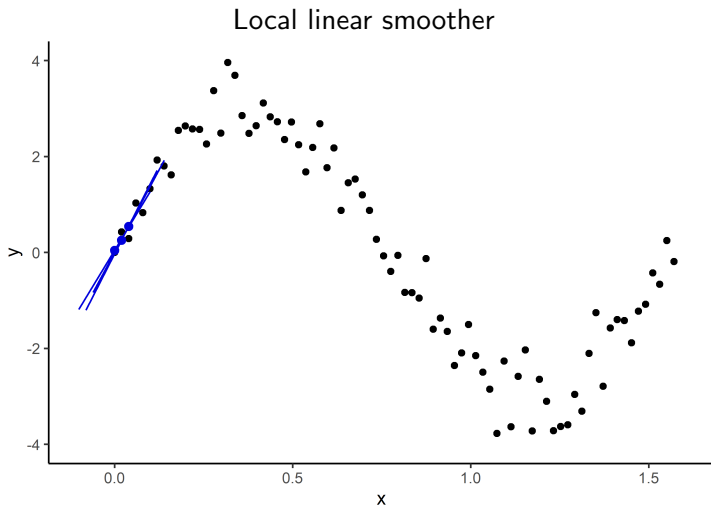
Kernel Smoothing



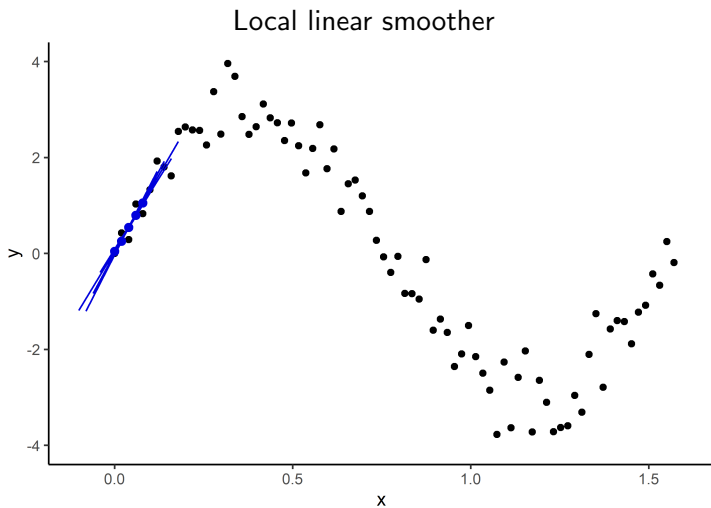
Kernel Smoothing



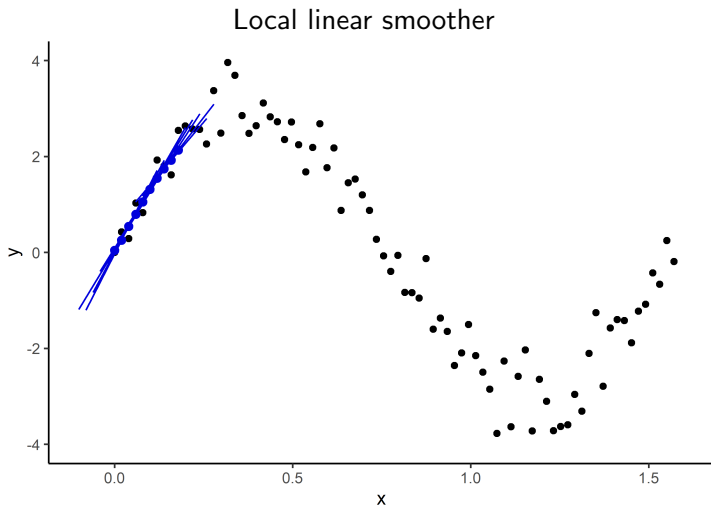
Kernel Smoothing



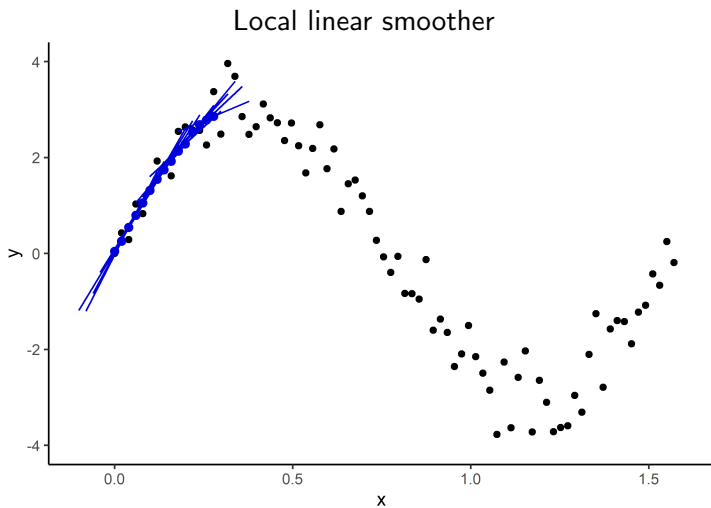
Kernel Smoothing



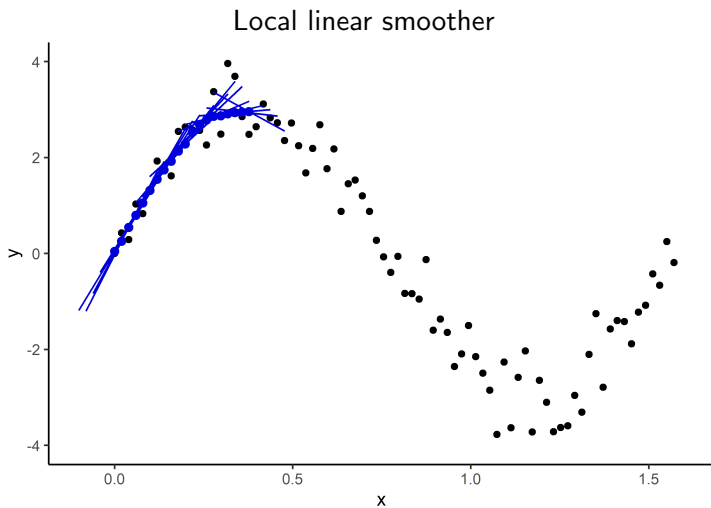
Kernel Smoothing



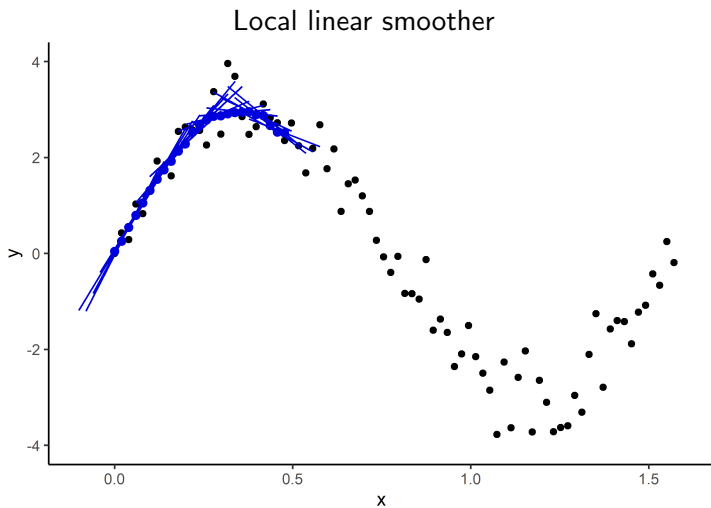
Kernel Smoothing



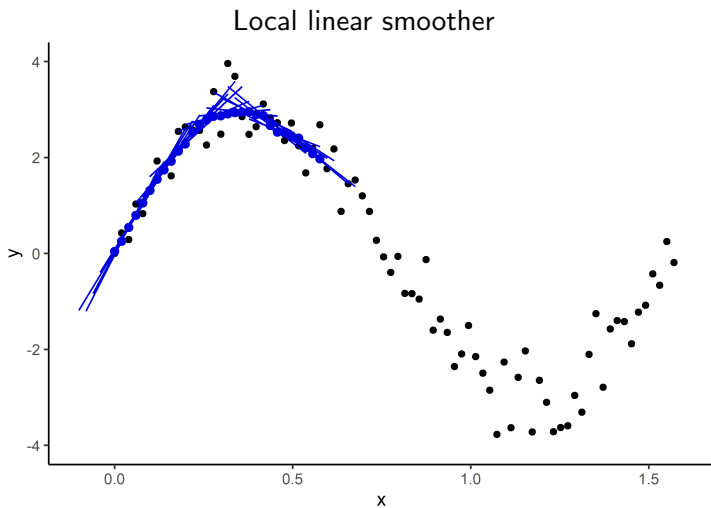
Kernel Smoothing



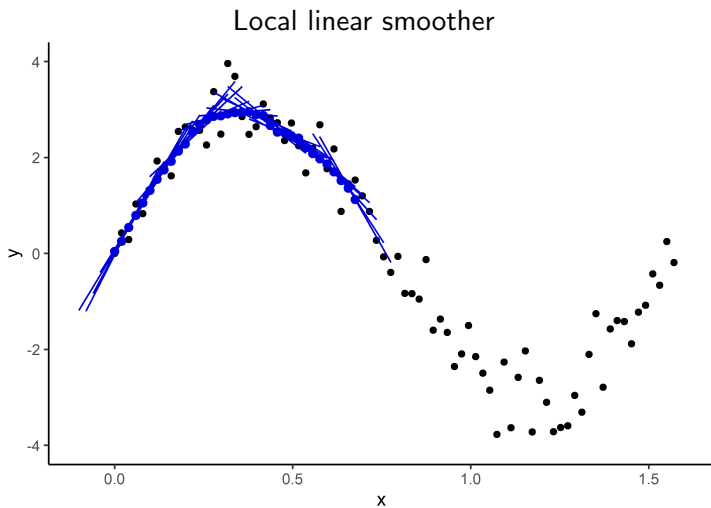
Kernel Smoothing



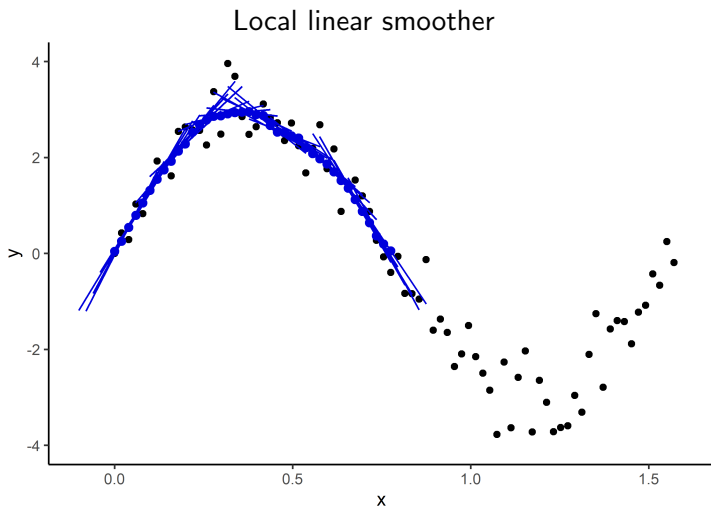
Kernel Smoothing



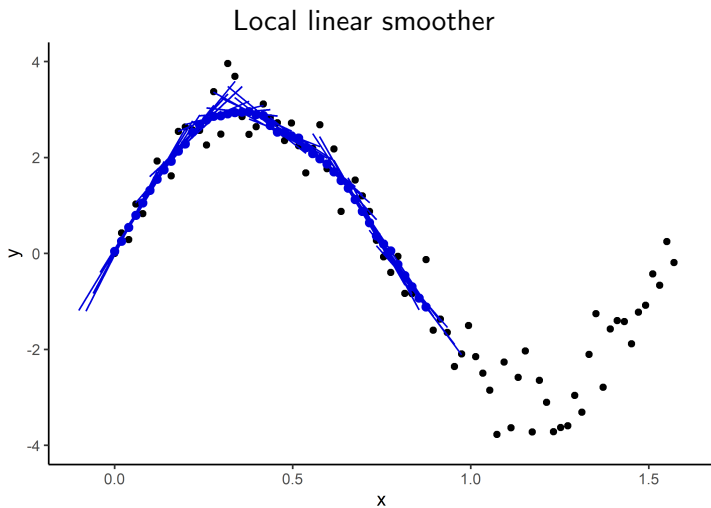
Kernel Smoothing



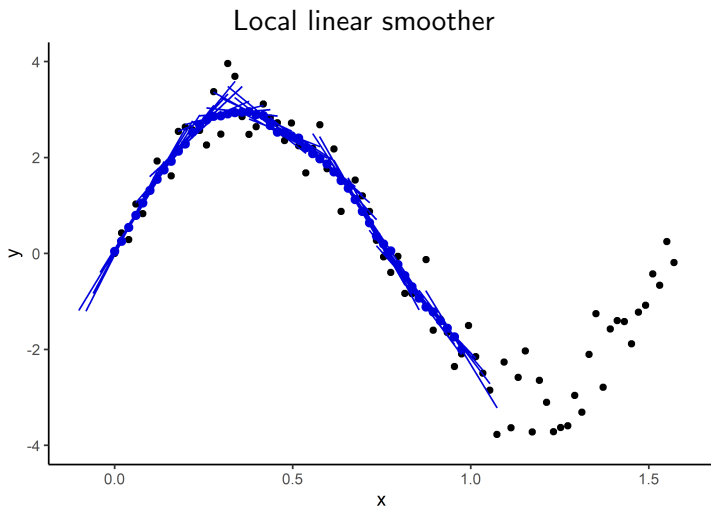
Kernel Smoothing



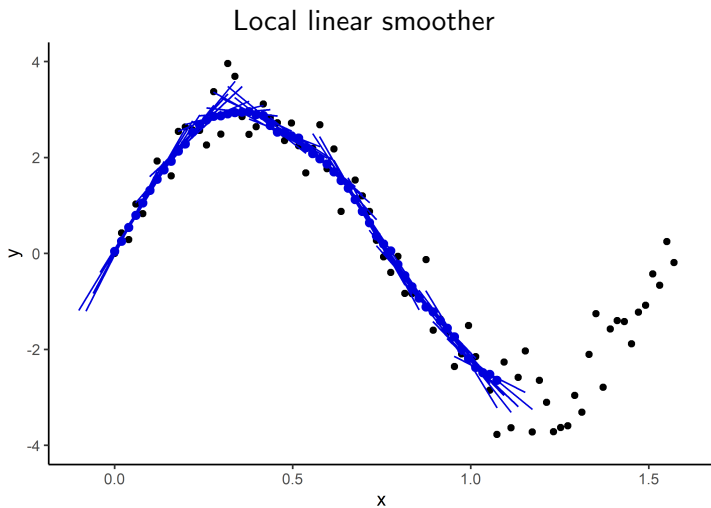
Kernel Smoothing



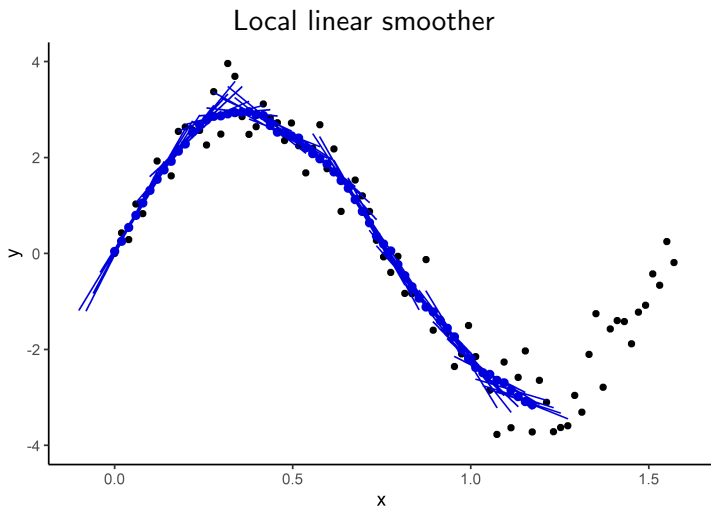
Kernel Smoothing



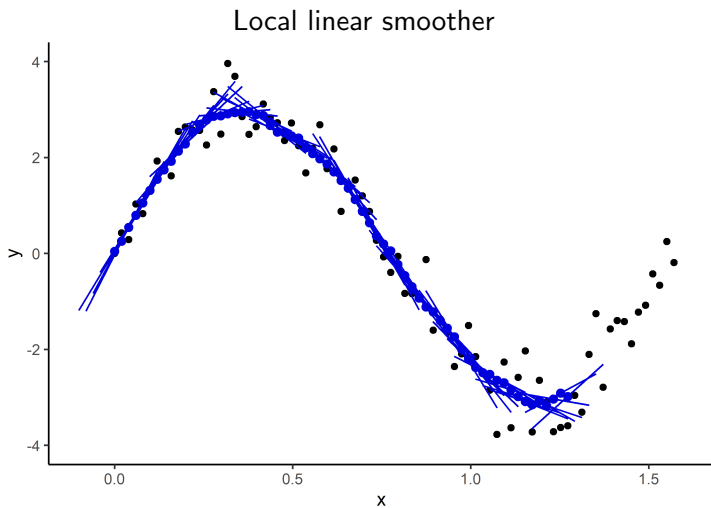
Kernel Smoothing

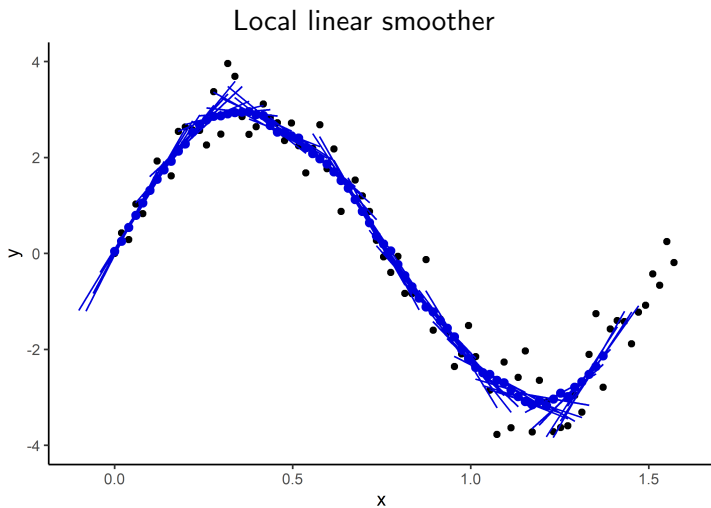


Kernel Smoothing

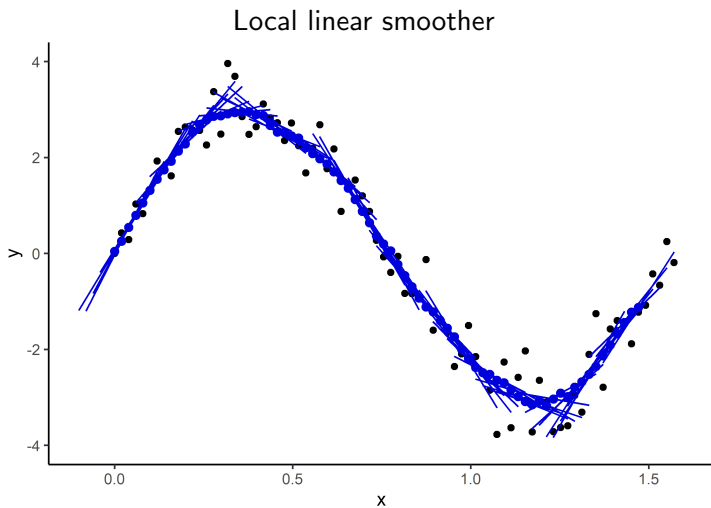


Kernel Smoothing

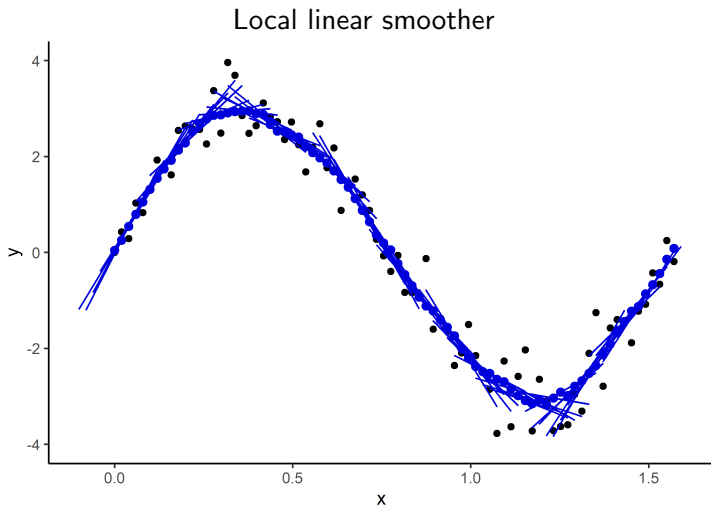




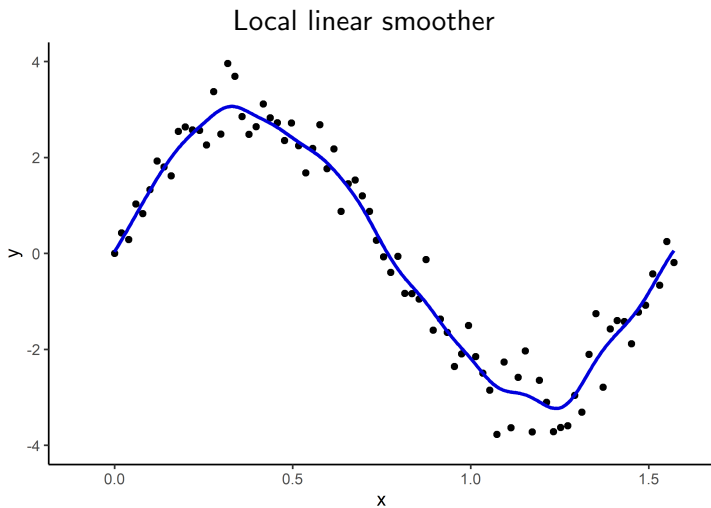
Kernel Smoothing

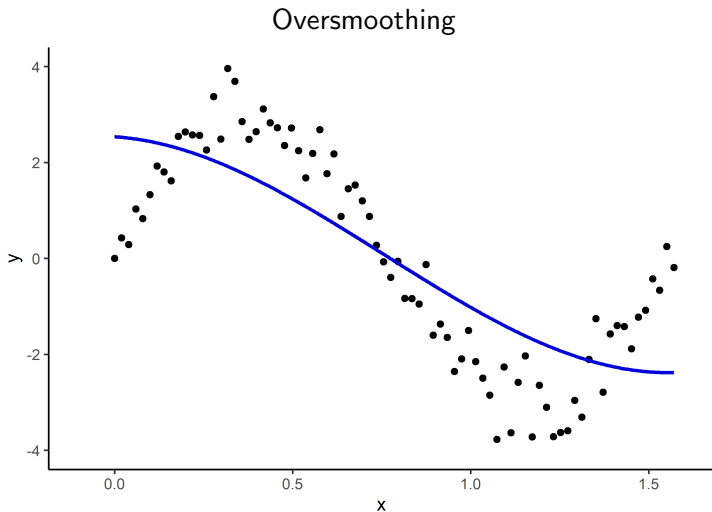


Kernel Smoothing

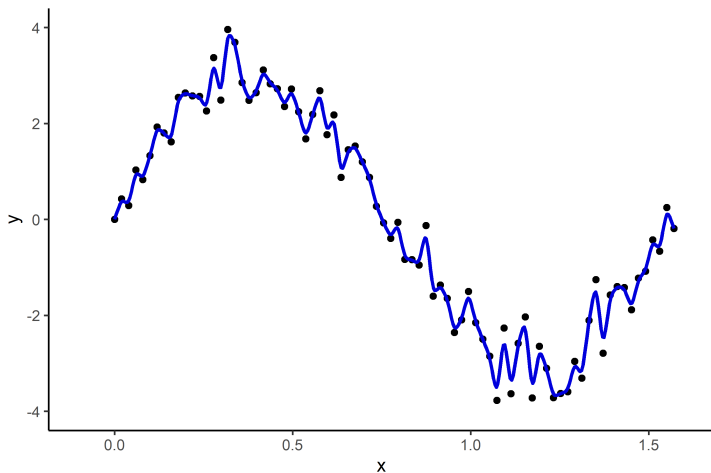


Kernel Smoothing





Undersmoothing



Mean function estimate

Local linear smoother with **global** bandwidth

$$\sum_{i=1}^n \sum_{j=1}^{N_i} \left[K \left(\frac{T_{ij} - t}{h} \right) Y_{ij} - \beta_0 - \beta_1(x - T_{ij}) \right]^2 \rightarrow \min$$

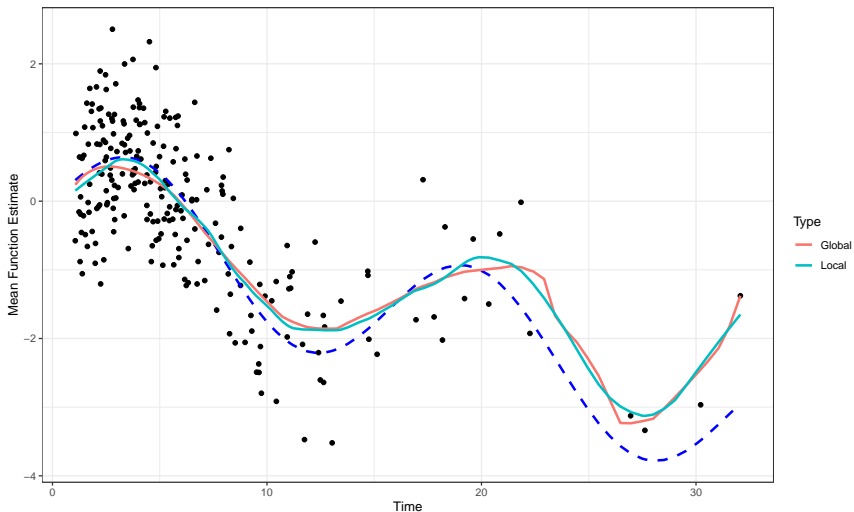
Local linear smoother with **local** bandwidth (Fan & Gijbels (1992))

$$\sum_{i=1}^n \sum_{j=1}^{N_i} \left[\alpha(T_{ij}) K \left(\frac{T_{ij} - t}{h} \alpha(T_{ij}) \right) Y_{ij} - \beta_0 - \beta_1(t - T_{ij}) \right]^2 \rightarrow \min$$

optimal $\alpha(\cdot) \sim f^{1/5}(\cdot)$

Kernel Smoothing

Simulation



Covariance function estimate

Local linear smoother with global bandwidth

$$\sum_{i=1}^n \sum_{\substack{j_1=1 \\ j_1 \neq j_2}}^{N_i} \sum_{j_2=1}^{N_i} \left[K \left(\frac{T_{ij_1} - s}{h}, \frac{T_{ij_2} - t}{h} \right) Z(T_{ij_1}, T_{ij_2}) - \beta_0 - \beta_{11}(s - T_{ij_1}) - \beta_{12}(t - T_{ij_2}) \right]^2 \rightarrow \min$$

- $\hat{c}(s, t) = \hat{\beta}_0(s, t)$
- **goal**: adapt the local bandwidth method

Covariance function estimate

Local linear smoother with local bandwidth

$$\sum_{i=1}^n \sum_{\substack{j_1=1 \\ j_1 \neq j_2}}^{N_i} \sum_{j_2=1}^{N_i} \left[\alpha(T_{ij_1}) \alpha(T_{ij_2}) K \left(\frac{T_{ij_1} - s}{h} \alpha(T_{ij_1}), \frac{T_{ij_2} - t}{h} \alpha(T_{ij_2}) \right) Z(T_{ij_1}, T_{ij_2}) \right. \\ \left. - \beta_0 - \beta_{11}(s - T_{ij_1}) - \beta_{12}(t - T_{ij_2}) \right]^2 \rightarrow \min$$

Known issues:

- symmetry of $\hat{c}(s, t)$ (OK for symmetric kernels)
- positive definiteness of $\hat{c}(s, t)$ (particularly depends on h)
- optimal $\alpha(\cdot)$
- optimal h

① Motor Oil Data

The dataset contains amount of Fe particles depending on operating time and a number of oil changes. Data were collected 2006 – 2016 from 29 heavy-duty army vehicles.

- Load the variable `df.motor` from the `motoroil.RData` file and plot it (see Figure 1).
- Use functions from the file `functionsM7777.R` to fill the data (see Figure 2).
- Try to neglect the number of oil changes and put all groups together (see Figure 3).
- Fill the data using one of mentioned methods (see Figure 4).
- Do the same with using the FPCA package and compare results (see Figures 5, 6).
- (optional) Is the number of oil changes negligible? Conduct the fANOVA analysis. Is it correct to do it?

Problems to solve

Motor Oil Data

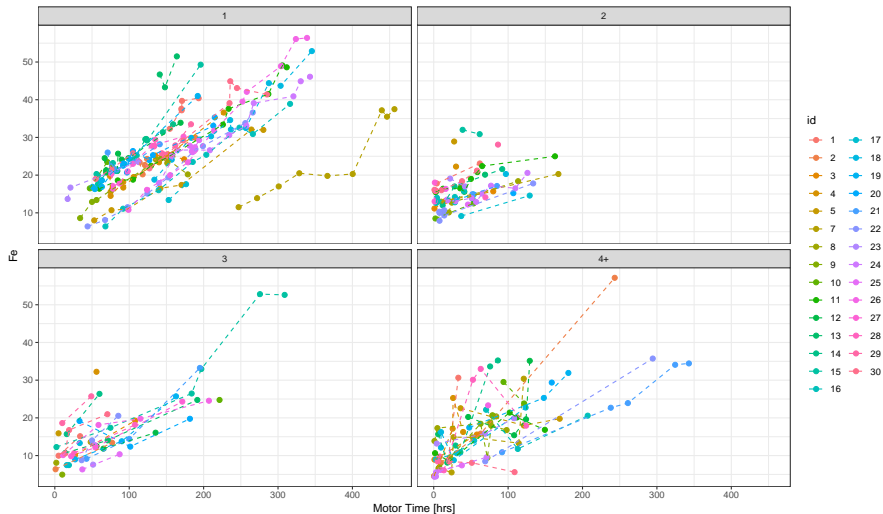


Figure 1.

Problems to solve

Motor Oil Data – filled

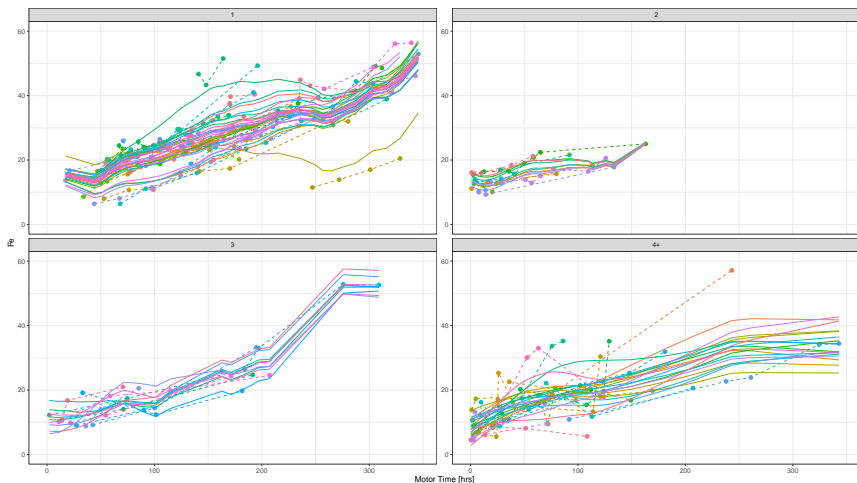


Figure 2.

Problems to solve

Motor Oil Data

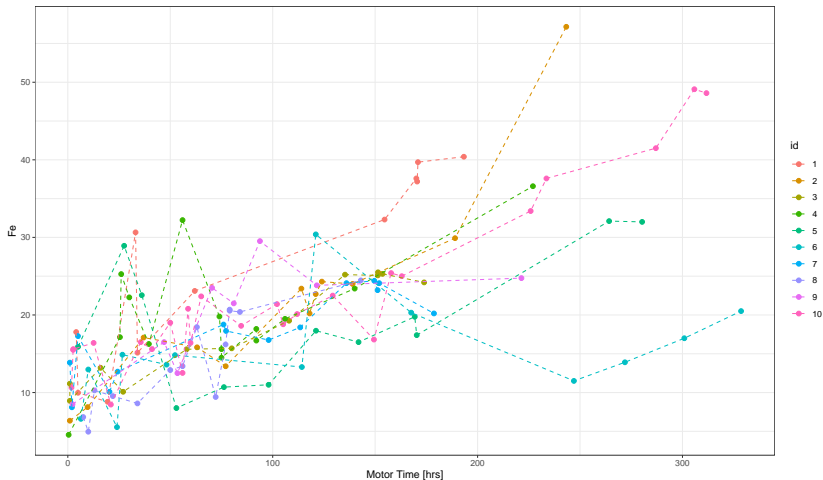


Figure 3.

Problems to solve

Motor Oil Data – filled

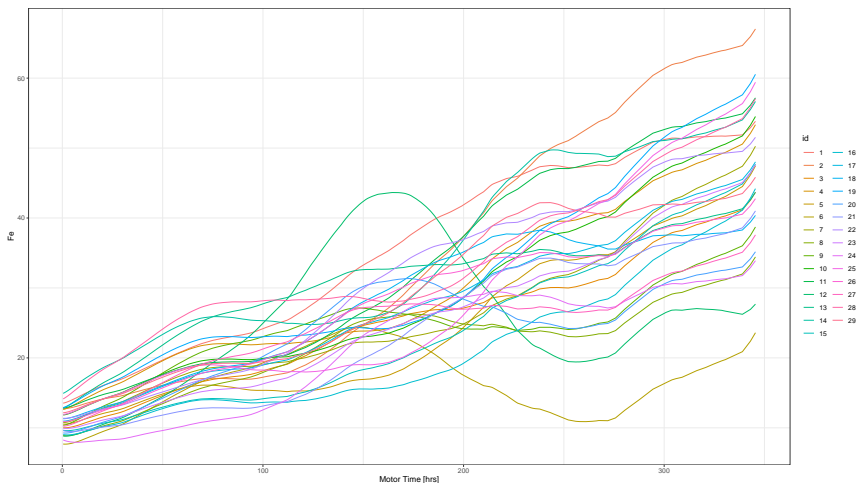


Figure 4.

functionsM777.R

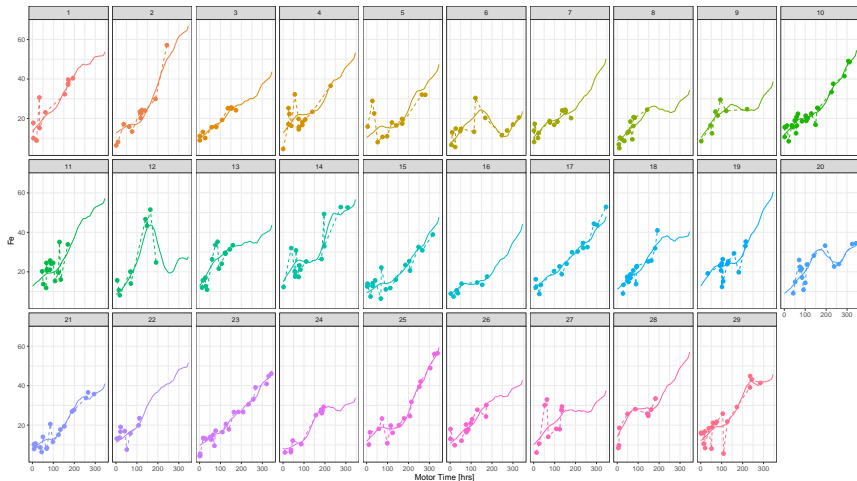


Figure 5.

FDAPACE

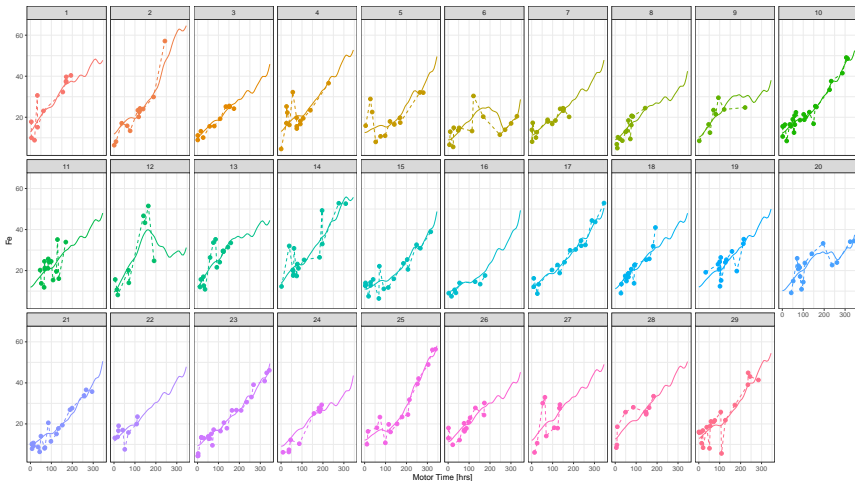


Figure 6.